# Benchmarking of Pay Components in CEO Compensation Design 

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#### Abstract

Eighty-nine percent of S\&P500 companies report benchmarking CEO pay components. Analyzing a panel of CEO compensation data entailing 1,251 S\&P 1500 firms during 2007-2013, we find that: 1) total compensation benchmarking less effectively explains CEO compensation than does component-of-pay benchmarking; 2) the strength of compensation components' adjustment to their benchmarks appears similar across the various components; the only exception is the salary that adjusts less pronouncedly to its benchmark, and 3) benchmarking is used not only regarding the level of compensation components but also regarding the weight of each component in total compensation. We discuss possible motivations for pay component benchmarking.


JEL classification: G34, G38, J31, J33, M12, M52

Keywords: Benchmarking, CEO compensation, peer group, target pay, pay structure, pay mix, pay components.

## 1. Introduction

In recent decades, senior executives' compensation has drawn intense academic and professional scrutiny. A central issue in these debates is the methodology employed by boards of directors and compensation committees to determine chief executive officer (CEO) pay.

In this study, we focus on the practice of compensation benchmarking, in which a given firm compares CEO compensation with the compensation packages of peer CEOs at comparable companies. Previous empirical research (Albuquerque, De Franco, \& Verdi, 2013; Bizjak, Lemmon, \& Nguyen, 2011; Bizjak, Lemmon, \& Naveen, 2008; Faulkender \& Yang, 2010; Laschever, 2013) has established that peer pay and benchmarking play an important role in determining total CEO compensation.

We extend the benchmarking research by analyzing the benchmarking of the components of CEO pay. Motivated by the description of benchmarking practices in compensation committee reports, we examine the following three questions: Is each pay component benchmarked separately and differently than other pay components? Is the structure of compensation (weight of each pay component in total pay) benchmarked as well? And, is pay component benchmarking a better description of benchmarking practices in US public firms than total pay benchmarking?

We employ two research strategies (and samples) to answer our research questions, and focus primarily on the benchmarking of three major pay components: Salary, equity-based compensation, and non-equity performance pay. First, we read the compensation-committee reports (Form DEF 14A) of S\&P 500 firms in fiscal year 2013, and document any statement referring to benchmarking CEO pay components. We find that approximately $89 \%$ of firms explicitly state that they benchmark at least
one pay component. Further, about $75 \%$ of firms declare that they benchmark all three major pay components. These figures indicate that these firms examine separately the distribution of salary, equity-based compensation, and non-equity-based compensation among peers to determine the level of each pay component to their CEO. We also examine whether companies target CEO compensation structure (weight of each pay component in total CEO compensation), and find that approximately $30 \%$ of firms explicitly declare in their proxy statement that they benchmark the compensation mix.

Our second empirical strategy employs detailed compensation data to examine the effectiveness of component benchmarking in explaining the cross-sectional and time-series variation in the reported compensation of CEOs. We analyze a relatively large CEO compensation database of 4,892 firm-year observations and 70,347 peeryear observations on 1,251 unique firms included in the $\mathrm{S} \& \mathrm{P}$ Composite 1500 during 2007-2013, and we consequently make three important observations. First, we find that component benchmarking describes the data, i.e., describes actual pay practice, more effectively than does total compensation benchmarking.

Second, when benchmarking pay components, the adjustments of CEO salary to that of selected peers are significantly less pronounced than the adjustments of the two other major pay components, i.e., non-equity performance pay and equity pay. We also identify an economically large yet statistically insignificant difference in the adjustment-to-peers coefficient between equity pay and non-equity performance pay.

Third and perhaps most novel, we present evidence supporting the contention that benchmarking is used not only when determining CEO total pay or pay component levels, but also when designing the structure of CEO pay, in terms of the proportions of the various pay components. Boards attend to the proportion of each component in total pay, and, according to a simple estimate, correct CEO pay so that it closes about
half of the previous-year gap in the weight of the pay component between the given CEO and their peer group. ${ }^{1}$

Our findings regarding the benchmarking of CEO pay components are unexpected. According to economic theory, total pay benchmarking helps firms provide competitive CEO pay packages that serve to retain valuable human capital (Holmstrom \& Kaplan 2003). However, competitive pay packages do not explicitly indicate benchmarking each pay component separately nor benchmarking the mix of CEO compensation.

We discuss potential motivations for pay component benchmarking. Several of these motivations are in line with optimal compensation design. For example, boards may rely on other firms' compensation design, as these designs provide information regarding the optimal compensation structure. Further, boards might have a difficult time setting competitive pay to their CEOs if the compensation structure deviates from the compensation structure of other firms, since each pay component has a different certainty equivalence. For example, the certainty equivalence of a stock option is lower than that of cash salary. Granting the CEO a similar pay structure (weight of each pay component in total pay) as that of her peers, alleviates concerns that the CEO compensation is not competitive.

We also note that benchmarking of pay components may arise from external players' involvement in the process of CEO compensation design (regulators, compensation consultants, proxy advisors, and even board of directors), and express the concern that such external interventions could sometimes lead to suboptimal

[^1]compensation. Examining the motivations behind pay component benchmarking and whether benchmarking leads to an optimal compensation design are important agendas for future research, yet they are beyond the scope of our primarily explorative study.

The rest of the paper is organized as follows. Section 2 presents the literature review and outlines our hypotheses. Section 3 describes the data and the sample selection process. Section 4 outlines our basic empirical model of the level of CEO pay components. Section 5 reports and discusses the benchmarking tests' results; Section 6 presents potential motivations for benchmarking pay components and pay structure; and Section 7 concludes.

## 2. The Benchmarking Practice

### 2.1. Background

A common practice in the CEO pay-setting process is comparing pay with that of peer groups. In a given case, the set of peer firms is selected by the directors and the compensation committee members, who often engage external compensation consultants. Outside the firm, proxy advisors also use peer compensation as their benchmark when evaluating senior executive pay plans. According to the benchmarking method, the current level of a given firm's CEO compensation is compared to that of a peer group of CEOs at similar firms, in which similarity is typically based on industry, size, and talent flow considerations (where talent flow embodies past sources and destinations of the firm's executives). In such analyses, pay below the median is usually considered "below market."

In 2006, to facilitate transparency, the Securities and Exchange Commission (SEC) adopted new proxy disclosure rules that require firms to report all companies in the peer group or survey when the use of peer groups has a material impact on executive
compensation. The disclosure must include the names of the individual companies and a detailed explanation for the selection of these companies as peers. The SEC's 2006 disclosure requirements enable researchers to examine the impact of actual peer group pay on the level of CEO compensation.

Existing studies have shown that the median CEO total pay in the peer group helps explain CEO pay. Further, the impact of median peer CEO pay on firm CEO pay exceeds the impact of stock market performance on pay (e.g., Faulkender \& Yang, 2010; Bizjak et al., 2011; Albuquerque et al., 2013) ${ }^{2}$.

### 2.2. The Debate over the Virtue of Benchmarking

The purpose of the benchmarking method is to adjust the level of executive compensation. A senior executive who is compensated improperly may potentially resign from the company or neglect her duties. In order to retain valuable human capital, the company should follow the market compensation standards. The benchmarking of CEO compensation is a practical and efficient mechanism to gauge the market wage (Holmstrom and Kaplan, 2003).

The main concern regarding benchmarking is that it may be used to justify pay raises unrelated to CEO or firm performance. Critics of the use of peer group benchmarking argue that powerful CEOs persuade compensation committees to select peer firms in a way that inflates CEO pay (O'Reilly, Main \& Crystal, 1988; Main, O’Reilly, \& Wade, 1995; Newman \&Mozes, 1999).

[^2]The pay inflation critique of benchmarking is supported by some studies. Faulkender and Yang (2010) show that the level of CEO compensation at a potential peer company affects its likelihood of being chosen as a compensation peer. This bias towards highly paid peers is particularly strong in firms where the peer group is smaller, where the CEO is also the chairman of the board of directors, where the CEO has longer tenure, and where directors are busy serving on multiple boards. Similarly, Bizjak et al. (2011) and Laschever (2013) show that when firms deviate from the economic model of peer firm choice, they tend to pick larger firms and firms with higher CEO pay. These biases in peer group selection are more evident in smaller, less visible firms where arguably management has more discretion in selecting the peer group. Last, Hayes and Schaefer (2009) provide an interesting explanation for picking highly paid peers. They develop a game-theoretic model of the "Lake Wobegon Effect", and show that boards may rationally inflate CEO pay to influence market perception regarding the quality of firm management and value of the firm.

However, the debate about the manipulation of benchmarking is ongoing and unresolved. Cadman and Carter (2014) do not find evidence for opportunistic selection of peers, while Larcker, McClure, and Zhu (2020) estimate that in a third of their firmyear observations, peers were selected opportunistically. Albuquerque et al. (2013) find that the CEO talent component of pay (approximated based on past abnormal performance, the size of the firms the CEO has managed in the past, and media coverage) is from two to ten times larger than the self-serving component of pay (captured by board structure, antitakeover provisions, and ownership concentration). Finally, in line with the "competition for managerial talent" view, Cremers and Grinstein (2014) report that benchmarking practice prevails largely in industries in
which new CEOs tend to come from outside the firm. In contrast, no evidence for benchmarking obtains in industries with few outside CEOs.

A recent addition to the literature indicates that the benchmarking method may also serve as a motivational tool. Francis, Hasan, Mani, and Ye (2016) find that firms with relatively high peer quality in terms of managerial skills exhibit performance superior to firms with relatively poor peer quality. Choosing a more skilled peer group can improve firm performance in two ways: one, CEOs may be motivated by highly ranked peers to increase their own work efforts; and/or, two, CEOs may learn successful new policies from these presumably excellent peers.

### 2.3. Existing Evidence on Benchmarking CEO's Total Pay

Benchmarking of total pay has been extensively studied in prior research. Bizjak et al. (2008) document that the use of peer groups is widespread. Of a random sample of 100 firms listed in the S\&P 500 index, 96 reported that peer groups had been used in determining compensation. Further, 73 firms mentioned targeting one or more of the components of pay at either the median or mean of the peer group. Bizjak et al. (2008) also find that CEOs who are paid below the median level of their industry- and sizematched peers receive increases in total pay that are $\$ 1.3$ million per year higher than the raises received by their counterparts whose pay is above the peer group median. In each sample year, approximately one-third of the executives with pay below their peer group's median receive pay adjustments that advance them to or above the median level of pay in their peer group. Another interesting finding is that the effect of peer group benchmarking on changes in CEO pay is stronger than is the effect of stock price performance on changes in CEO pay.

Recent literature also offers several estimates of the sensitivity of CEO pay to peers' pay. Faulkender and Yang (2010), Albuquerque et al. (2013), and Laschever (2013), estimate an elasticity of CEO pay with respect to median peer CEO pay of 0.38 to 0.58 . Bizjak et al. (2011) find that the annual increase in compensation closes about one-third of the difference in pay between a given CEO and their peer group median.

### 2.4. Hypotheses

### 2.4.1. Benchmarking pay components

Previous studies have focused on the benchmarking of CEO's total pay. When we review DEF 14A forms, however, we find that most firms state that they benchmark each component of total pay. When determining CEO pay, many firms examine each pay component separately by comparing it to the median value, or to a specified percentile, of this pay component among their chosen peers.

The following examples are useful. The 3M Company states in its 2012 proxy statement that "the Committee generally aims to provide the Company's executives whose performance meets the Company's expectations with Total Cash Compensation that is at or very close to the median of the corresponding compensation paid to executives in the benchmarking groups, and with long-term incentive compensation delivered through annual grants having values that are close to the average of the $50^{\text {th }}$ and $75^{\text {th }}$ percentiles of the corresponding grant values provided to executives in the benchmarking groups". Another example is the 2014 proxy statement of Align Technology, where the company disclosed a target for base salary ( $50^{\text {th }}$ percentile), a target for total cash compensation $\left(65^{\text {th }}\right.$ to $75^{\text {th }}$ percentile), and a target for equity compensation ( $50^{\text {th }}$ to $75^{\text {th }}$ percentile)

Based on the examples above, we propose that each of the CEO pay components has unique character, purpose, and behavior. Consequently, each of the main pay components may be benchmarked separately and potentially differently. CEO compensation packages comprise six main components of pay: (i) base salary; (ii) bonus; (iii) non-equity incentive plan compensation; (iv) option grants; (v) restricted stock grants; and (vi) other pay. Our empirical work considers each of the six components separately. However, our central empirical tests divide total compensation into the three components we consider most significant: (i) salary; (ii) non-equity performance pay (i.e., bonuses and non-equity incentive plan compensation); and (iii) equity pay (i.e., restricted stock and option grants). This division of compensation enables us to differentiate between performance-based components and non-performance-based components as well as between cash-based components and equitybased components.

Our primary hypothesis is as follows

Hypothesis 1: When the level of CEO pay component $X$ deviates from the norms at peer firms in year t -1, the given CEO will sustain an adjustment towards the peer pay correction in year t pay component X .

Hypothesis 1 examines the previously untested prediction that pay components are benchmarked. In addition, we will examine whether or not all components of pay are benchmarked to the same extent, i.e., employ the same adjustment parameters.

### 2.4.2. Benchmarking the pay mix

We notice in DEF 14 forms that many firms also mention benchmarking peer compensation structure (mix of pay components). For example, Northeast Utilities declares in its 2014 proxy statement that "We target the mix of compensation for our

Chief Executive Officer and the other Named Executive Officers so that the percentages of each compensation element are approximately equal to the competitive median market mix."

Motivated by this description, in our empirical work, we posit that benchmarking is employed with attention to the structure of CEO pay as well. Prior studies demonstrate the importance of well-designed pay structure. Mehran (1995) argues that proper pay structure may motivate managers to increase firm value and finds that firm performance is positively related to the percentage of executive compensation that is equity-based.

The structure of CEO compensation has changed considerably over time. Pay mix changes arguably reflect (i) increases in firm size and business complexity that perhaps increase potential agency problems; (ii) a better understanding over time of what attracts, motivates, and retains CEOs; and (iii) regulatory changes. The history of CEO pay practices is reviewed in Murphy (2013) and in Edmans, Gabaix, and Jenter (2017).

Pay structure is important because it determines the relative magnitude of the various incentives aligned to the different components of total pay. To illustrate, equity pay, for example, may incentivize a given CEO's attention to firm stock price, while non-equity performance pay may incentivize the CEO's efforts in other directions including accounting profitability, long-term efficiency, survival, or executive suite coordination (e.g., Guay, Kepler, \& Tsui, 2019). In such a case, the balance between equity and non-equity performance pay in the compensation plan signals to the CEO the relative importance of pursuing a high stock price. Boards and compensation consultants are likely to seek an optimal balance between different pay components, and one solution is to follow the pay structure of peer firms.

However, benchmarking pay structure is not a straightforward conclusion. Rationally, different companies should design a mix of compensation components to suit their particular needs, firm strategy, and CEO characteristics (including CEO age, equity holdings, and other characteristics). Nevertheless, learning from peer firms may provide insights into options for optimal compensation structure. This is especially true when the chosen peer is a relatively successful firm, and this partly explains compensation committees' tendency to choose comparatively highly paid CEOs as peers. To the extent that peer firms are comparable and successful, learning from the selected peers' pay design, including pay mix, is both relevant and prudent for the firm in question. We discuss this argument in more detail in Section 6.

If benchmarking is also employed with respect to a given pay component's share in total CEO pay, we posit the following:

Hypothesis 2: Benchmarking the pay structure affects CEO pay components; CEOs whose share of pay component $X$ in total compensation is below (above) the peer group median in year t-1, will receive an upward (downward) adjustment in the level of pay component X in year t .

## 3. Samples and Data

We construct two data sets. The first includes DEF 14A forms of all S\&P 500 firms for fiscal year 2013. This sample enables us to examine company statements concerning CEO pay benchmarking policy. The second data set comprises detailed CEO compensation data for all S\&P Composite 1500 firms and their compensation peers in the years 2006-2013. This data set enables us to examine the actual CEO compensation practices of these firms. For brevity we denote the first sample as the "policy sample" and the second sample as the "CEO pay practice sample."

### 3.1. The Policy Sample

We review DEF 14A forms of S\&P 500 firms for fiscal year 2013 to find information on benchmarking of CEO pay components, benchmarking of CEO total compensation, and benchmarking of CEO compensation structure (mix of pay components). Our analysis focuses on three main pay components: salary; non-equity performance pay; and equity pay.

First, we search the Compensation Discussion and Analysis (CD\&A) section of the DEF 14A forms for information on benchmarking CEO total pay and the pay components. This information can be found in the chapters describing the executive compensation philosophy and objectives, the pay setting process, components of pay, and peer groups. We use the following keywords: median; $50^{\text {th }}$; mid-point; percentile; component; element; peer; benchmark; comparator; competitive; and market practice.

Second, we search the DEF 14A forms for explicit statements indicating that firms employ peer group data to determine the mix of CEO pay components. This information can be found in the CD\&A chapters describing executive compensation philosophy and objectives, peer groups, and the pay mix. We use the following keywords: mix; structure; proportion; and weight.

Before proceeding, we note that for 24 of the 505 firms in our policy sample, we do not find any DEF 14A forms. In addition, four firms use vague statements regarding benchmarking, e.g., a statement that they may consult national compensation surveys, hence we include them in the missing information total count.

Table 1 summarizes our findings. About 75\% of firms state that they benchmark all three pay components, and an additional $14 \%$ explicitly mention that they benchmark one or two of our three main pay components. Twenty-four firms report
benchmarking total compensation only, and 4 firms do not mention benchmarking at all in their DEF 14A forms. Thus, in summary, 449 out of the 505 firms- $89 \%$-use some form of pay component benchmarking.

In 383 ( $85.3 \%$ ) out of the 449 firms benchmarking pay components, firms benchmark their pay components to the median of the peer group or state that they use peer groups to set the CEO individual pay components without specifying any target percentile. More interestingly, 66 firms-approximately $13 \%$ of the policy sampleexplicitly set pay component targets different from the median of their peers.
[Insert Table 1 about here]

Regarding benchmarking CEO total compensation, 66.5\% (336) of the firms report benchmarking total CEO pay in addition to benchmarking pay components; an additional $4.8 \%$ (24) state they target total compensation only. Interestingly, the fraction of firms declaring total compensation benchmarking is lower than the fraction declaring pay component benchmarking.

Finally, in 154 (30.5\%) of the 505 firms, we find statements that the firm also employs the peer group to determine the mix between the various components of CEO pay. This explicit reference to the structure of pay benchmarking supports our novel Hypothesis 2, which posits that CEO pay structure is also benchmarked.

### 3.2. The CEO Pay Practice Sample

The initial CEO pay practice sample comprises 10,481 firm-year observations on S\&P Composite 1500 index firms during 2007-2013. The S\&P Composite 1500 combines three indices: the S\&P 500, the S\&P MidCap 400, and the S\&P SmallCap 600. In December 2006 the SEC introduced new amendments requiring firms to disclose their peer group when the use of peer groups is material in the pay setting
process. Accordingly, peer group data have become available in definitive proxy statements (DEF 14A) beginning in fiscal year 2006. However, the SEC's 2006 disclosure rules modify also the compensation reporting format and redefine some of the compensation components. This makes 2005 and 2006 compensation component data not fully comparable. Since benchmarking requires comparison to previous year data, the first year we examine is 2007.

We focus on executives classified as CEOs by Standard \& Poor's ExecuComp database from which we collect CEO compensation data. We exclude 93 firm-year observations with no available compensation data for the current and/or previous year, and 35 observations of zero values for total compensation. Pay component observations with negative values are defined as missing values. We further drop 2,000 firm-year observations of CEOs who were replaced or appointed during the current or previous year to exclude partial compensation or exceptional high payments (e.g., golden parachutes, severance pay, golden handshakes, and sign-on bonuses). This reduces sample size to 8,353 firm-year observations.

For each of the 8,353 firm-years, we find the disclosed list of compensation peers. We assemble this information from two sources. Peer information for 2007 and 2008 is based on manually collected data from the Compensation Discussion and Analysis (CD\&A) section of the proxy statements. ${ }^{3}$ These data are described in Albuquerque, De Franco and Verdi (2013). We construct peer lists for 1,639 firm-year observations based on Albuquerque et al. (2013)'s dataset.

Peer data for 2009-2013 are collected from the Executive Compensation Analytics (ECA) database provided by Institutional Shareholder Services (ISS). The

[^3]ECA database starts in 2008, however, we preferred to use the dataset of Albuquerque et al. (2013) for the overlapping year. We construct peer lists for 4,884 firm-year observations using the ECA database. Together with the list based on Albuquerque et al. (2013), we amass an explicit list of peers for 6,523 firm-years, which is $78 \%$ of our initial list. The remaining $22 \%$ firm-year observations are firms that are missing in the ECA database.

Next, given our list of peer CEOs, we seek current and prior year compensation data for 112,328 peer firm-year observations (peer-year observations, in short). Thus, there are on average more than 17 peers for each CEO. We find available compensation data on ExecuComp and ECA for 98,432 peer-year observations because among the peers there are foreign and private firms that are not included in the ExecuComp or ECA databases.

Given available peer and peer compensation data, we exclude 33 disclosing firm-years that report only one or two peers. In addition, we exclude 396 firm-years with missing compensation data for $50 \%$ or more of their chosen peers. ${ }^{4}$ Finally, we exclude 34 observations of co-CEOs and 1,168 observations of firms in the financial services industry (industries 45-48 in the Fama-French industry classification). Table 2 summarizes the exclusion process. The final sample comprises 4,892 firm-year observations, representing 1,251 unique disclosing firms.
[Insert Table 2 about here]

We collect data on peer-based pay targets from the ECA database when available. The ECA database reports firms' target percentile for three pay measures: salary; bonuses and non-equity incentive plan compensation; and total compensation.

[^4]For each disclosing firm and compensation component, we calculate the peer-based target pay level based on peer compensation data and the target pay percentile set by the firm's board. Three conventions are used in calculating the target pay level. First, when the ECA data set does not specify any target percentile, we assume the target equals the median pay of the peers. Second, because ECA does not specify distinct targets for equity pay, we use the total compensation percentile targets to calculate the target pay levels for equity pay, stock awards, and option awards. Third and finally, in cases where the target is expressed as a range, e.g., $50-60 \%$, of peers' pay, we use the middle of the target range (55\%).

Nonmedian targets are relatively scarce. In approximately $12 \%$ of the firm years in our final sample (589 out of the 4,892 firm-years) there are nonmedian compensation targets. This statistic corresponds closely with our finding in Table 1, showing that 13\% of the S\&P 500 firms employ a nonmedian target for at least one of the pay components. Notably, $4 \%$ of those nonmedian targets ( 23 firm years) are lower than the median, $92 \%$ are in the third quartile (between 0.5 and 0.75 ), and $4 \%$ exceed 0.75 .

Regarding the structure of compensation, we assume that the benchmark for the weight of pay component X in total compensation is its median proportion among the peers.

We use two procedures to mitigate the potential effect of outliers in the highly skewed compensation data. First, as common in the compensation literature, all compensation data are winsorized at the $2.5 \%$ and $97.5 \%$ levels within each year. Changes in pay are also winsorized at these percentages. Second, we use the logarithm of the pay measures and the logarithmic change of pay. The logarithmic transformation is common, and it facilitates comparison with previous studies.

Stock return data are from the Center for Research in Security Prices (CRSP) database. Data on other financial variables (sales, ROA, market-to-book ratio, and financial leverage) that have been found in prior research to explain variations in CEO pay, are extracted from Standard \& Poor's Compustat database. Data on the CEO's name, age, and possible dual role as CEO and Chairman of the board are collected from the Execucomp database.

### 3.3. Descriptive Statistics for Our CEO Pay Sample

Our final pay data sample comprises 4,892 firm-year CEO pay observations in S\&P 1500 firms during 2007-2013. These observations are distributed almost uniformly across the sample years. Every firm-year observation also includes information concerning the peers. The average (median) peer group for our sample firms includes about 18 (16) firms; when we deduct missing peer compensation data, the average (median) number of peers with available compensation data per firm is 14 (13). The final sample includes 70,347 peer-year observations. The average and median number of peers are consistent with prior studies (e.g.; Faulkender \&Yang, 2012; Albuquerque, Franco, \& Verdi, 2013).

Table 3 offers extensive descriptive statistics of annual CEO pay, change in annual pay and pay structure, all compared to peer firms' CEOs. Table 3 Panel A focuses on the pay levels at our disclosing firms (for brevity, denoted hereafter as firms). The average (median) annual CEO total compensation is 6.964 (5.260) million dollars. The mean (median) sum of salary and discretionary bonus is 994 (850) thousand dollars, which is very close to the one million dollar cap on deductible compensation. The average (median) stock awards of 2.343 (1.500) million dollars is almost double the average (median) annual level of option awards. The mean non-equity incentive compensation is 1.296 million dollars. The "other pay" component, which is the sum
of the change in pension value, non-qualified deferred compensation earnings, and all other compensation, is shown for completeness, and its mean is 778 thousand dollars.

Target pay measures based on previous year peer data are presented to the right of the CEO pay statistics. Target pay statistics resemble those of the actual CEO pay. For example, the average (median) of the peer-based target total compensation is 6.850 (5.617) million dollars while the average (median) total compensation of CEOs is 6.964 (5.260) million dollars.

Table 3 Panel B provides descriptive statistics for various pay change measures, estimated as the first difference in $\operatorname{Ln}($ pay $)$. The variation in the number of observations is due to our excluding from the analysis those cases in which the pay component takes the value of zero in either the current or previous year. The average yearly change in total compensation over the sample period is $8.5 \%$, and is mainly due to an average annual increase of $9.6 \%$ in equity pay comprising option and stock awards. Other pay components such as salary and non-equity incentive plan compensation increase at an average annual rate of $4 \%$ over the sample period.

Changes in pay targets based on peer group data are reported on the right-hand side of Table 3 Panel B. The average change in the total compensation target is 5.8\%, which is smaller than the corresponding $8.5 \%$ average increase in total compensation among disclosing firms. However, differences between the medians are less pronounced, especially when we compare the components of pay. Also noteworthy is the mean negative change of $3.8 \%$ in peer "other pay." It warns us that "other pay" may be a residual term that is perhaps less well-controlled or benchmarked. In the main analysis of this study we ignore "other pay."

Table 3 Panel C presents descriptive statistics on the compensation structure. Pay components are reported as a percentage of total compensation. Note that median ratios do not sum to $100 \%$ because the sum of the medians is not the median of the sum. Also, the sum of the mean ratios in Table 3 Panel C is lower by 3 percentage points than $100 \%$ because pay components are positively skewed and each pay measure is winsorized at the $2.5 \%$ and $97.5 \%$ levels.

On average, nearly $19 \%$ of total CEO compensation is in salary, $2 \%$ in discretionary bonuses, $19 \%$ in non-equity incentive compensation, $31 \%$ in stock awards, $17 \%$ in option awards, and $9 \%$ in "other pay." These statistics show that U.S. CEOs receive most of their pay in the form of performance-sensitive compensation. The right-hand side of Table 3 Panel C reports statistics concerning the target ratios of various pay measures to total compensation, based on peers' compensation data. The compensation structure of the sample firms is similar to that of the peer firms, as consistent with Hypothesis 2 positing that pay structure is also benchmarked.

## [Insert Table 3 about here]

Table 4 documents summary statistics for characteristics and explanatory variables common in the CEO compensation literature. The mean (median) annual stock return is $15 \%$ ( $12 \%$ ). The sample period includes both the Great Recession (years with negative stock returns) and the following years of recovery. The mean and the median return on assets (ROA) is 0.05 . We employ one-year lagged annual sales revenue as our proxy for firm size. The average lagged sales is $\$ 8,135$ million, almost four times the median sales of $\$ 2,138$ million. This is because, as is usual, the distribution of firm sales is skewed towards large values. The average firm risk, represented by the standard deviation of monthly stock returns in the 36 months preceding the end of the previous fiscal year, has a mean (median) of $0.11(0.10)$. The
one-year lagged market-to-book (MTB) ratio, a measure of a given firm's growth opportunities, is defined as the ratio of the market value of equity to the book value of equity at the end of year $\mathrm{t}-1$. The mean (median) lagged MTB is 1.82 (1.51). One-year lagged financial leverage is measured as book value of total liabilities divided by the sum of the book value of total liabilities and the market value of equity, all of which are measured at the end of year t -1. One-year lagged leverage has a mean (median) of 0.35 (0.33). CEO age is a proxy of CEO's work experience. The average CEO age is 56 years. The youngest CEO is 29 years old, while the oldest is 93 . Finally, only $16 \%$ of the CEOs in the sample serve also as chairman of the board.

## [Insert Table 4 about here]

## 4. Evidence on the Determinants of CEO Pay and Its Components

The conventional specifications of existing models of CEO total compensation (Albuquerque et al., 2013; Bizjak et al., 2008; Faulkender \& Yang, 2012, 2010; Laschever, 2013) is
(1) $\operatorname{Ln}\left(\right.$ CEO compensation $\left._{\mathrm{i}, \mathrm{t}}\right)$

$$
\left.\begin{array}{l}
=\alpha_{0}+\alpha_{1} \operatorname{Ln}\left(\text { Sales }_{\mathrm{i}, \mathrm{t}-1}\right)+\alpha_{2}\left(\text { Stock return }_{\mathrm{i}, \mathrm{t}}\right)+\alpha_{3}\left(\text { Stock return }_{\mathrm{i}, \mathrm{t}-1}\right) \\
+\alpha_{4}\left(\operatorname{ROA}_{\mathrm{i}, \mathrm{t}}\right)+\alpha_{5}\left(\operatorname{ROA}_{\mathrm{i}, \mathrm{t}-1}\right)+\alpha_{6} \operatorname{Ln}\left(\text { Risk }_{\mathrm{i}, \mathrm{t}-1}\right)+\alpha_{7}\left(\operatorname{MTB}_{\mathrm{i}, \mathrm{t}-1}\right) \\
+\alpha_{8}\left(\text { Leverage }_{\mathrm{i}, \mathrm{t}-1}\right)+\alpha_{9}\left(\text { CEO Age }_{\mathrm{i}, \mathrm{t}}\right)+\alpha_{10}(\text { CEO Duality Dum } \\
\mathrm{i}, \mathrm{t}
\end{array}\right)
$$

where i indexes firms and t indexes year. We fit an analogous model for the following six different pay components: (i) salary; (ii) performance pay-the sum of bonus, option awards, stock awards, and non-equity incentive plan compensation; (iii) nonequity performance pay-the sum of bonus and non-equity incentive plan
compensation; (iv) equity pay-the sum of option awards and stock awards; (v) stock awards; and (vi) option awards.

The explanatory variables in Equation 1 include the following firm and CEO characteristics: (i) the natural logarithm of sales in the previous year (a measure of firm size); (ii) stock returns and returns on assets (ROA) in years t and $\mathrm{t}-1$ (firm's performance indicators); (iii) the natural logarithm of the standard deviation of the monthly stock return in the 36 months preceding the end of the previous fiscal year (represents firm's risk); (iv) lagged market-to-book (MTB) ratio (a proxy for growth opportunities); (v) lagged leverage; (vi) CEO age; and (vii) CEO duality (a dummy variable that equals 1 when the CEO also serves as Chairman). We further include dummy variables for each unique combination of industry and year. Industry classification is based on the 49-industry classification of Fama and French (1997). Finally, the residual $e_{i t}$ is a firm-year specific error term that is assumed to be correlated within firms and heteroskedastic. As such, we cluster standard errors at the firm level. The traditional compensation model of Equation 1 is our baseline model in the rest of the paper.

Table 5 reports the results of estimating Equation 1 for total pay and for six different pay components, as described above. All seven regressions are statistically significant at least at the $1 \%$ level, and their adjusted R -squares range from 0.38 to 0.62 . Consistent with existing evidence on the relation between CEO compensation and firm size (Tosi et al., 2000; Gabaix, Landier, \& Sauvagnat, 2014; Edmans et al., 2017), the coefficient of $\operatorname{Ln}$ (Sales it,-1 ) is positive and highly significant for all compensation components. The positive coefficient of firm size likely indicates that the managerial talent and skills needed for running larger and more complex firms are scarce and therefore command higher compensation

The coefficients of stock return and lagged stock return are positive and statistically significant for all pay components except salary. This illustrates that CEO is rewarded (punished) for good (poor) stock performance. The pay performance relation is documented and widely studied in previous research (see, e.g., Edmans et al., 2017).

However, the relation between CEO compensation and accounting performance is unstable and cryptic, as current and lagged ROA coefficients are in general of opposite sign. Among pay components, only the total elasticity of non-equity performance pay to ROA (the sum of the coefficients of ROA and one-year lagged ROA) is on average positive, perhaps because non-equity performance pay is typically linked to accounting performance metrics. Interestingly, equity pay-especially option awards-are negatively related to ROA. This finding may reflect a cross-sectional phenomenon in which firms with relatively higher ROAs prefer to grant compensation in the form of non-equity performance pay, traditionally linked to accounting performance, rather than in the form of option awards. This result is consistent with the negligible effect of ROA on total CEO compensation documented in previous studies (e.g., Faulkender \&Yang, 2010, 2012; Laschever, 2013).

According to the principal-agent theory, riskier firms have to pay more than do less risky firms to induce their CEOs to exert greater effort to maximize firm value, and must also pay more to compensate their CEOs for the greater wealth uncertainty (Cheng, Hong, \& Scheinkman, 2015). Inconsistent with this prediction, however, we find a negative relation between firm risk and CEO pay. Previous studies have also shown this negative relation (see, e.g., Cohen \& Lauterbach, 2008; Faulkender \& Yang, 2012) Our finding is also in line with Albuquerque et al. (2020), who examine the
relation between compensation and risk, and conclude that CEOs with riskier pay packages do not receive an economically large extra compensation.

The coefficient on the lagged market to book ratio is significantly positive for both equity pay and option awards, and negative yet insignificant for base salary. Generally, monitoring CEOs in high growth (high market-to-book) firms is more challenging, given that these CEOs make many uncertain future-dependent investment decisions. Accordingly, growth firms use more equity-based compensation to incentivize their CEOs to exert effort and make investment decisions that maximize firm value.

As for the other explanatory factors, lagged leverage has a negative and significant effect on equity pay, which indicates that banks and debtholders' monitoring cuts equity pay. CEO age has a significant influence on salary and non-equity performance pay, probably reflecting the effect of work experience. Further, as consistent with previous studies (e.g., Faulkender \& Yang, 2010, 2012), we find that CEOs who serve also as chairmen of their boards receive higher pay.

## [Insert Table 5 about here]

Overall, the results are consistent with findings in earlier studies of CEO total compensation. However, we identify some variations in the explanatory power of the various factors across pay components. The salary, equity pay, and non-equity performance pay components each exhibits distinctive features. ${ }^{5}$ Therefore, examining each pay component separately, as we do in this study, enhances understanding of the determinants of CEO pay.

[^5]
## 5. Evidence on Benchmarking in the Level and Structure of CEO Pay

### 5.1. Univariate Evidence on Benchmarking in CEO Compensation

The benchmarking hypothesis (Hypothesis 1) predicts that CEOs with belowtarget pay in year $\mathrm{t}-1$ receive a pay raise in year t that is higher than the respective pay raise of CEOs who earn above-target pay in year $\mathrm{t}-1$. We begin by comparing the changes in pay of those CEOs who earn above and below the target pay level, where the target pay is based on peer compensation. Table 6 Panel A documents that the mean and median logarithmic changes in pay for those CEOs paid below the target are higher than the respective changes in pay for those CEOs paid above the target in the prior year. These preliminary univariate test results are consistent with Hypothesis 2.

Both parametric and nonparametric tests indicate that the pay raise gaps between the below- and above-target groups are statistically different from zero at the $1 \%$ level for all compensation components. We use $t$-statistics when comparing mean changes, and Wilcoxon rank sum $z$-statistics for the nonparametric tests. Also significant is that the number of observations in the 'below target' groups is generally larger than the number of observations in the 'above target' groups. This difference indicates firms' tendency to select highly paid CEOs as their peers (Faulkender \& Yang, 2010; Bizjak et al., 2011).

Among pay components, salary exhibits not only the lowest year-to-year change, but also the lowest difference between the 'below target' and 'above target' groups. For the other pay components, the mean (median) proportional pay raise gaps between below- and above-target paid CEOs range from 0.25 to 0.36 ( 0.07 to 0.21 ). Among the various pay measures examined, performance pay exhibits the widest average pay gap: a $25 \%$ mean increase for below-target paid CEOs versus an $11 \%$ pay
cut for above-target paid CEOs. CEOs who earn above the peer group in the previous year receive a pay cut in the following year with the salary component the only exception. These pay cuts in the 'above target' group appear to challenge the popular view of powerful CEOs determining their own pay.

## [Insert Table 6 about here]

Overall, the results in Table 6 Panel A indicate that on average CEOs receive larger (lower) raises when their pay is below (above) their target pay, when target pay is assessed on the basis of peer-group pay. Importantly, this result holds for all nine pay and pay component specifications reviewed in Table 6 Panel A, as well as when we measure the raise in dollar terms. For example, the average (median) total pay increase of below target firms is about 1067 (513) thousand dollars, significantly higher than the average (median) total pay increase of above target firms that is about 706 (210) thousand dollars. In sum, univariate tests appear to strongly support the benchmarking of the level of pay components (Hypothesis 1).

We next examine whether benchmarking is also employed in determining the structure of CEO pay. Because each component of pay may induce a different effort scheme from the CEO, an optimal pay mix is essential. Our Hypothesis 2 contends that boards turn to comparable firms to gauge the optimal composition of CEO pay.

In order to examine Hypothesis 2, we compute the average year-by-year changes in the weight of various pay components in total compensation for two groups: (i) CEOs whose previous-year weight of pay component X in total compensation is above the previous year peer group median; and (ii) CEOs whose previous-year weight of pay component X in total compensation is below the peer group median in the previous year

Table 6 Panel B documents the results. The mean change in the weight of each pay measure in total compensation is positive for the "below median" group and negative for the "above median" group. Thus, as predicted by Hypothesis 2, on average, the weight of pay component X in firm i is corrected towards the peer group median weight of component X . Further, $t$-tests indicate that for all pay components, the difference between the mean weight change of above- and below-median firms is statistically significant at the $1 \%$ level. These results provide preliminary evidence that firms benchmark not only CEO pay levels but also CEO pay structure.

To further examine whether benchmarking affects the structure of pay, we estimate the following regression:

$$
\begin{aligned}
& \text { (2) } \Delta\left(\frac{\text { pay component } \mathrm{X}}{\text { CEO total compensation }}\right)_{i, t}=\beta_{0} \\
& +\beta_{1}\left[\left(\frac{\text { peer pay component } \mathrm{X}}{\text { peer total compensation }}\right)_{i, t-1}-\left(\frac{\text { pay component } \mathrm{X}}{\text { CEO total compensation }}\right)_{i, t-1}\right] \\
& +\beta_{2}\left(\text { IndustryDum }_{\mathrm{i}, \mathrm{t}}\right) \times\left(\text { YearDum }_{\mathrm{t}}\right)+\varepsilon_{\mathrm{i}, \mathrm{t}}
\end{aligned}
$$

where i indexes firms and t indexes time. The dependent variable is the change in the weight of pay component X in total compensation from year $\mathrm{t}-1$ to year t . The independent variable that captures benchmarking is the difference between the median weight of pay component $X$ in peer firms' total compensation and the corresponding weight at a disclosing firm, both at year $\mathrm{t}-1$. We refer to this difference as the distance from peer group median in year t -1. If the distance from peer group median is positive (negative), the weight of pay measure X in total compensation in firm i is below (above) the corresponding peer group median. If boards try to adjust their CEO pay structure according to peer group median structure, the coefficient $\beta_{1}$ in Equation 2 should be positive. We further include industry-year fixed effects and cluster standard errors at the firm level.

The results, reported in Table 6 Panel C, show that the coefficients of the distance from the peer group median weight in year $\mathrm{t}-1$ are positive and statistically significant at $1 \%$ level at least, for all pay components. Moreover, interestingly, the coefficient estimates in Table 6 Panel C hover at approximately 0.5 , indicating that about half of the pay weight gaps are closed within a year. Evidently, pay structure corrections are not minute.

### 5.2. Adding Compensation Benchmarking to the CEO Pay Model

Benchmarking focuses on the changes in CEO pay. For this reason, we start by differencing the pay component version of our baseline CEO pay model in Equation 1 and obtain:
(3) $\Delta \operatorname{Ln}\left(\right.$ CEO compensation component $\left.\mathrm{X}_{\mathrm{i}, \mathrm{t}}\right)$
$=\beta_{0}+\beta_{1} \Delta \operatorname{Ln}\left(\right.$ Sales $\left._{\mathrm{i}, \mathrm{t}-1}\right)+\beta_{2} \Delta\left(\right.$ Stock return $\left._{\mathrm{i}, \mathrm{t}}\right)+\beta_{3} \Delta\left(\right.$ Stock return $\left._{\mathrm{i}, \mathrm{t}-1}\right)$
$+\beta_{4} \Delta\left(\right.$ ROA $\left._{\mathrm{i}, \mathrm{t}}\right)+\beta_{5} \Delta\left(\mathrm{ROA}_{\mathrm{i}, \mathrm{t}-1}\right)+\beta_{6} \Delta \operatorname{Ln}\left(\operatorname{Risk}_{\mathrm{i}, \mathrm{t}-1}\right)+\beta_{7} \Delta\left(\mathrm{MTB}_{\mathrm{i}, \mathrm{t}-1}\right)$
$+\beta_{8} \Delta\left(\right.$ Leverage $\left._{i, t-1}\right)+\beta_{9}\left(\right.$ IndustryDum $\left._{\mathrm{i}, \mathrm{t}}\right) \times\left(\right.$ YearDum $\left._{\mathrm{t}}\right)+\varepsilon_{\mathrm{i}, \mathrm{t}}$,

Next, we add to the above equation our benchmarking measures. Bizjak et al. (2011) focus on total CEO pay and define the benchmarking explanatory variable as the natural logarithm of the peer CEO-based pay divided by firm CEO pay level, both at year t -1. The implicit assumption is that compensation committees and board members try to correct the previous year distortions (relative to peers) in their total CEO pay. We use an analogous definition for each pay component, i.e., our first benchmarking explanatory variable for pay component X is the ratio of the peer-based target for pay component X (for example, the median of pay component X among the peers) to the actual level of firm CEO pay component X , both at the previous year.

The second benchmarking variable is novel in the literature and relates to pay structure benchmarking. Our univariate tests in the previous section support the pay structure benchmarking idea (Hypothesis 2). Benchmarking pay structure necessarily affects the level of the pay components. For example, if the weight of pay component X in total compensation is below the peer group median, its adjustment towards the median peer weight requires an increase in the level of pay component X that is separate and supplementary to the other required adjustments of the level of X . The benchmarking variable representing the pay structure gap is the difference between the median weight of pay component X in total compensation among the chosen peers and the corresponding weight for a sample firm CEO.

Because the benchmarking of the pay component level must be integrated with the benchmarking of the pay structure, the resulting proposed model for each pay component is:
(4) $\Delta \operatorname{Ln}\left(\right.$ CEO compensation component $\left.X_{i, t}\right)=\beta_{0}+$ $\beta_{1} \operatorname{Ln}\left(\right.$ Relative compensation component $\left.X_{i, t-1}\right)+$ $\beta_{2}\left[\left(\frac{\text { peer pay component } \mathrm{X}}{\text { peer total compensation }}\right)_{i, t-1}-\left(\frac{\text { pay component } \mathrm{X}}{\text { CEO total compensation }}\right)_{i, t-1}\right]+\beta_{3} \Delta \operatorname{Ln}\left(\right.$ Sales $\left._{i, t-1}\right)+$ $\beta_{4} \Delta\left(\right.$ Stock return $\left._{\mathrm{i}, \mathrm{t}}\right)+\beta_{5} \Delta\left(\right.$ Stock return $\left._{\mathrm{i}, \mathrm{t}-1}\right)+\beta_{6} \Delta\left(\right.$ ROA $\left._{\mathrm{i}, \mathrm{t}}\right)+\beta_{7} \Delta\left(\right.$ ROA $\left._{\mathrm{i}, \mathrm{t}-1}\right)+$ $\beta_{8} \Delta \operatorname{Ln}\left(\right.$ Risk $\left._{\mathrm{i}, \mathrm{t}-1}\right)+\beta_{9} \Delta\left(\right.$ MTB $\left._{\mathrm{i}, \mathrm{t}-1}\right)+\beta_{10} \Delta\left(\right.$ Leverage $\left._{\mathrm{i}, \mathrm{t}-1}\right)+$ $\beta_{11}\left(\right.$ IndustryDum $\left._{\mathrm{i}, \mathrm{t}}\right) \times\left(\right.$ YearDum $\left._{\mathrm{t}}\right)+\varepsilon_{\mathrm{i}, \mathrm{t}}$,
where i indexes firms, X indexes the compensation components, t indexes year, and all other explanatory variables are as previously defined (see also the Appendix).

Table 7 presents the results of fitting the integrated benchmarking model to the data. For brevity, only statistically significant coefficients at the $10 \%$ level or higher
are shown. The estimated coefficients of $\operatorname{Ln}($ relative compensation component $X$ ) are positive and highly statistically significant for all pay components. The magnitude of the coefficients ranges from approximately 0.07 for salary to 0.32 for performance pay. Thus, a CEO with performance pay $1 \%$ below (above) the target pay in year $\mathrm{t}-1$ receives, ceteris paribus, a performance pay increase in year t that is $0.32 \%$ larger (smaller) than that of a CEO whose year t-1 performance pay equals the target performance pay. The adjustment coefficients of the various compensation components in Table 7 tend to be slightly lower than the adjustment coefficient of 0.31 estimated by Bizjak et al. (2011) for total pay using data for 2006. However, the clear conclusion remains that the gap in CEO pay component X relative to peers triggers a significant revision (i.e., correction towards the peers) in the next year. The results also document that boards only partially adjust CEO pay, which indicates that boards use benchmarking cautiously.

The coefficients of our pay structure benchmarking variable, the distance from peer group median in the weight of pay component X in total compensation, are positive and significant at $1 \%$ level for all pay components. The coefficient estimates of the weight difference range from a low of 0.027 for salary to a high of 0.47 for option awards. This coefficient signifies, for example, that a CEO whose proportion of equity pay is $1 \%$ below (above) the peer group median receives, ceteris paribus, an increase in equity pay that is about $0.45 \%$ larger (smaller) than a CEO whose proportion of equity pay is similar to the peer group median. The coefficient on the distance from peer group median in the salary regression is the lowest across all pay components, implying an incremental increase (decrease) in salary pay of only $0.03 \%$ for a CEO whose proportion of salary in total pay is $1 \%$ below (above) the peer group median.
[Insert Table 7 about here]

The results in Table 7 support Hypothesis 2 and show that year-by-year changes in CEO pay component X are also affected by the wedge between the firm and its peer group in the weight of pay component X in total compensation. Table 7 also confirms the effects of the relative pay status of the CEO in year t-1 (Hypothesis 1). In summary, CEO pay components are plausibly adjusted based on two benchmarking criteria: (i) the relative-to-peers' level of the pay component in the previous year; and (ii) the previous year difference in pay component weight in total compensation from its peer group median weight.

Whether the various pay components are benchmarked differently merits investigation. To test the difference across various pay components in the benchmarking coefficients, we employ the seemingly unrelated regressions (SUR) methodology. Two motivations exist for using the SUR framework: (i) to achieve more efficient estimation by combing information from different equations; and (ii) to test restrictions involving parameters from all equations. When examining different pay components, the residuals of these regressions should be correlated, due to the probability of common unobserved factors that influence year-by-year changes in pay across all pay components.

The equation system we use comprises the three major pay components: salary; non-equity incentive; and equity pay. For each component we use the model specified in Equation 4 above. For example, for salary we use:
(5) $\Delta \operatorname{Ln}\left(\right.$ Salary $\left._{\mathrm{i}, \mathrm{t}}\right)$

$$
\begin{aligned}
& =\beta_{0}+\beta_{1} \operatorname{Ln}\left(\text { Relative salary }_{\mathrm{i}, \mathrm{t}-1}\right) \\
& +\beta_{2}\left[\left(\frac{\text { peer's salary }}{\text { peer total compensation }}\right)_{i, t-1}-\left(\frac{\text { CEO's salary }_{\text {CEO total compensation }}^{i, t-1}}{}\right]\right. \\
& +\sum_{m=3}^{10} \beta_{\mathrm{m}} \operatorname{Controls}_{m, i}+\beta_{11}\left(\text { IndustryDum }_{\mathrm{i}, \mathrm{t}}\right) \times\left(\text { YearDum }_{\mathrm{t}}\right)+\varepsilon_{\mathrm{i}, \mathrm{t}}
\end{aligned}
$$

The null hypothesis is that the coefficients describing benchmarking are equal across the three pay components' regressions. The alternative hypotheses propose differences in the adjustment coefficients. For example, regarding $\beta_{1}$, we examine two alternatives

$$
\begin{aligned}
& \text { H1a }: \beta_{1, \text { Salary }} \neq \beta_{1, \text { non-equity performance pay }} \neq \beta_{1, \text { equity pay }} \text {, and } \\
& \text { H1b }: \beta_{1, \text { non-equity performance pay }} \neq \beta_{1, \text { equity pay }}
\end{aligned}
$$

The results of the SUR estimation are presented in Table 8 Panel A. The coefficients of the SUR estimation are consistent with those reported in Table 7. The differences in the magnitude of the estimated coefficients between Tables 7 and 8 are probably due to the decrease in the number of observations in the SUR analysis excluding firm-years without observations for all three major pay components.

Table 8 Panel B summarizes the results of F-tests examining cross-components (i.e., cross-equations) differences in the benchmarking coefficients. We find a significant difference in the coefficients when all three pay components are compared. The source of this variation is the clearly weaker benchmarking of CEO's salary.

One explanation for the less pronounced adjustment of the salary component is that the salary compensation is subject to the one million dollar tax deductibility rule. Thus, firms approaching the one million dollars' cap from below may find adjusting
their CEO base salary at the same rate as other components of pay to be more costly. Consequently, pay adjustment of salary becomes less pronounced.

To further explore the one million dollar cap explanation, we generate a dummy variable that equals 1 when CEO's previous year base salary is below 900 thousand dollars, and zero otherwise. Presumably, CEOs with a salary below 900 thousand dollars are less restricted by the one million dollar salary cap regulation. This dummy variable is then interacted with the benchmarking measure, $\operatorname{Ln}\left(\right.$ Relative salary $\left._{i, t-1}\right)$. Adding this interaction term to the SUR system, we find that for CEOs who earn a salary below 900 thousand dollars, the adjustment coefficient is 0.10 . This coefficient is statistically significantly higher than the over 900 thousand dollars respective coefficient of 0.065 . Thus, the one million dollar cap appears to mitigate salary adjustments.

However, the salary cap is a partial explanation only. The adjustment coefficient for the subsample of below 900 thousand dollar salary, 0.10 , is still markedly lower than the average adjustment coefficient of 0.21 estimated for the two other pay components (equity pay and non-equity performance pay). We conclude that the one million dollar cap cannot adequately explain the considerably lower magnitude of adjustment of the salary component documented in Table 8. The fact that salary is a "sure" cash pay component may cause boards to adjust it more conservatively than the two other uncertain and performance-related pay components.

Although the differences in adjustment coefficients between equity pay and non-equity performance pay are statistically insignificant, they are not trivial. For example, the coefficient on the weight difference of pay component $X$ from peer group's median is equal to 0.54 for equity pay and 0.37 for non-equity performance pay. This finding may indicate differences between equity and non-equity performance
pay that we do not have sufficient statistical power to confirm. Interestingly, some significant differences are found in our robustness tests, reviewed later in Section 5.4.
[Insert Table 8 about here]

Overall, the results in this section demonstrate that the practice of compensation benchmarking significantly affects year by year changes in CEO pay. These findings are consistent with our hypotheses and extend previous evidence on benchmarking.

### 5.3. Are Pay Components Benchmarked Differently from Total Pay?

This study analyses each pay component separately. However, it is possible to argue that only total compensation is benchmarked, and the pay component levels are consequential, i.e., adjusted later according to their proportion in total pay. We can directly test the proposition that pay components are benchmarked independently from total compensation by fitting the following model:
(6) $\Delta \operatorname{Ln}\left(\right.$ Pay component $\left.X_{i, t}\right)$

$$
\begin{aligned}
& =\beta_{0}+\beta_{1 \mathrm{~T}} \operatorname{Ln}\left(\text { Relative total pay } \mathrm{i}_{\mathrm{i}, \mathrm{t}-1}\right)+\beta_{1 \mathrm{X}} \operatorname{Ln}\left(\text { Relative pay component } \mathrm{X}_{\mathrm{i}, \mathrm{t}-1}\right) \\
& +\sum_{m=2}^{9} \beta_{\mathrm{m}} \text { Controls }_{m, i}++\beta_{10}\left(\text { IndustryDum }_{\mathrm{i}, \mathrm{t}}\right) \times\left(\text { YearDum }_{\mathrm{t}}\right)+\varepsilon_{\mathrm{i}, \mathrm{t}}
\end{aligned}
$$

Equation 6 allows both total pay and individual component benchmarking. It uses both previous-year relative total pay and previous-year relative pay component as explanatory variables. Under a conservative null hypothesis (benchmarking all pay components identically to total pay), the coefficient $\beta_{1 \mathrm{X}}$ in Equation 6 should equal 0 for all pay components. This is because according to the null there is only one set of benchmarking criteria, those based on total pay, i.e., the component-specific benchmarking criteria is redundant.

Further, in practice, to avoid distortions due to the multicollinearity, we regress relative pay component X on relative total pay, and use the residual of this regression instead of relative pay component X when fitting Equation 6. This pre-test orthogonalization process makes rejections of the null even more difficult.

Table 9 presents the results of fitting Equation 6 to the three key pay components data using the SUR (Seemingly Unrelated Regressions) methodology. First, as a baseline, we report results of a set of regressions with only the total pay benchmarking variable, and a set of regressions with only individual pay-component benchmarking. Then, we report results of regressions with both total pay and individual pay-component benchmarking variables.

In Table 9 regressions, the coefficients of both the total pay and pay component benchmarking variables are positive and statistically significant at the $1 \%$ level at least. However, the system-weighted $\mathrm{R}^{2}$ of the formulation that includes pay component benchmarking alone, 0.281 , is remarkably higher than the system-weighted $\mathrm{R}^{2}$ of the formulation with only total pay benchmarking (0.219). Further, when both pay component benchmarking and total pay benchmarking are used (in the third set of regressions reported in Table 9), the system weighted $\mathrm{R}^{2}$ improves only slightly-from 0.281 to 0.282 -relative to the set of regressions containing pay components alone. Pay component benchmarking alone seems to adequately explain the revision in pay component X, i.e., total pay benchmarking appears secondary in the pay component setting process.

## [Insert Table 9 about here]

More formal tests of the importance of pay component benchmarking can be conducted. The null hypothesis proposes that in our system of three pay components
regressions that includes both individual pay component and total pay benchmarking explanatory variables:

$$
\beta_{1, \text { salary }}=\beta_{1, \text { non-equity performance pay }}=\beta_{1, \text { equity pay }}=0 \text {. }
$$

These implications of the null are tested and rejected by the data at the $1 \%$ level. Clearly, the actual benchmarking of a pay component is affected not only by the benchmarking of total pay, but each pay component also receives special attention, i.e., benchmarking, of its own. In addition to comparing total CEO pay to that of peers, boards also compare CEO pay component X to that of peers.

We also run a set of regressions with total pay and mix of pay benchmarking as explanatory variables. These regressions essentially replace the pay component benchmarking terms on the right-hand side of Equation 6 with our pay structure benchmarking variables, the deviations of the pay component weight from its peers' median weight. In unreported results, we find that the coefficients of all total pay benchmarking variables and all pay structure benchmarking variables are highly statistically significant, and the system weighted $R^{2}$ is 0.270 , which is considerably higher than that of the set of regressions containing total pay benchmarking variables only. Benchmarking of the pay mix is demonstrably also employed by the compensation committee and the board.

In this context, we examine the last set of pay component regressions in Table 9. This last set is copied from Table 8 for comparison convenience; it combines pay component benchmarking with pay structure benchmarking. Notably, its system weighted $\mathrm{R}^{2}, 0.285$, is the highest in Table 9 , implying that assuming benchmarking of
individual pay components and benchmarking of the pay mix are the most successful explanation for CEO compensation. ${ }^{6}$

The tests in this section supplement our main tests and reinforce our conclusion that separate benchmarking attention is devoted to each key CEO pay component. Further, our evidence may also be interpreted to indicate that CEO compensation design is a bottom-up process, which builds from the individual pay components to total compensation.

### 5.4. Some Robustness Tests

We conduct several robustness tests on our reported results to remove noise and assist inference. First, we exclude those firm-year observations with a nonmedian target for pay components in order to observe the intensity of benchmarking in "regular" firms, i.e., in order to examine whether our results are due to "activist" firms with nonmedian targets. Reestimating Equation 4, excluding firms with nonmedian targets, yields almost identical coefficients. Our conclusions remain intact.

A second concern is that firms that do not use all available pay components to compensate their CEOs potentially introduce noise. To address this concern, we repeat the analysis confining the sample to observations where all pay components are nonzero. Overall, stronger adjustments to the target, i.e., larger coefficients, are found for all our benchmarking variables. Interestingly, in the SUR analysis, besides the difference between salary and other pay components (previously reported in Table 8), we also identify in this subsample significant differences between the equity pay and non-equity performance pay components. Non-equity performance pay adjusts more

[^6]vigorously to its relative pay level measure (the previous year ratio of firm non-equity performance pay to that of its peers) whereas equity pay responds more strongly to a "distortion," (relative to peer CEOs) in its weight in total compensation. These findings indicate that in the benchmarking of non-equity performance pay, the absolute dollar pay of peers is relatively heavily scaled, while in the benchmarking of equity pay the structure of compensation (weight of equity in total compensation) is relatively heavily scaled.

We further divide our sample into two subperiods, 2007-2009 and 2010-2013. In addition to examining the temporal stability of the benchmarking coefficients, this test also explores the effect of the 2009 revisions in compensation reporting regulations. ${ }^{7}$ Results are similar in the two subperiods, although the $\operatorname{Ln}\left(\right.$ Relative pay component $\left._{i, t-1}\right)$ coefficients appear somewhat lower in the second subperiod. ${ }^{8}$ We conclude that the benchmarking behavior of our sample firms appears fairly stable over time.

We also rerun the analysis using firm fixed effects instead of industry-year fixed effects to control for possible individual firm pay patterns. The coefficients of $L n($ relative pay component $X$ ) tend to increase, while the coefficients on the distance in the proportion of pay component X from peer group's median almost doubled compared to Table 7. The average adjusted $\mathrm{R}^{2}$ across regressions employing firm fixed effect is

[^7]0.25 as compared to 0.21 in the industry-year fixed effect version of Table 7. Evidently, the firm fixed effects formulation supports our findings and conclusions.

Another potential concern entails the compensation component in year $t$ being awarded using a multiyear compensation plan. To monitor the effect of such multiyear grants we add the lagged (year t-1) level of the examined pay component to the list of explanatory variables in our firm fixed effects formulation of the pay component change equation; this methodology is proposed in Bizjak et al. (2011). We find that all estimated coefficients of the benchmarking variables remain statistically significant and are even larger in absolute values compared to their counterparts in Table 7. In all the new pay component regressions, the coefficient of the lagged compensation variable is significantly negative and adjusted $\mathrm{R}^{2} \mathrm{~s}$ are higher than those reported in Table 7.

Further, benchmarking of non-equity performance pay is arguably more complex. According to this argument, when setting the target and criteria for non-equity performance pay, firms do not have previous-year data on peers' non-equity performance pay, due to previous-year payouts being typically based on accounting numbers that are determined after the audit of the financial results of year $t-1$. We reestimate the non-equity performance pay regression after replacing the benchmarking variable $\operatorname{Ln}\left(\right.$ relative pay component $\left.X_{i, t-1}\right)$ with $\operatorname{Ln}\left(\right.$ relative pay component $\left.X_{i, t-2}\right)$. The coefficient of $\operatorname{Ln}\left(\right.$ relative pay component $\left.X_{i, t-2}\right)$ is statistically insignificant.

We also replicate the main tests using two-digit SIC codes instead of the Fama and French (1997) 49-industry classification and find similar results. Finally, we reestimate all specifications with winsorized control variables (sales, stock return, ROA, standard deviation, market-to-book, and leverage) at $2.5 \%$ and $97.5 \%$ levels. The results are similar, and all conclusions remain intact.

## 6. Potential Motivations for Pay Component and Pay Structure Benchmarking

Our findings indicate a marked tendency by firms to benchmark each component of CEO compensation separately. Why do firms choose separate benchmarking rather than simply benchmarking total compensation? Explanations for benchmarking in the extant literature have focused on the role of total pay benchmarking in retaining valuable human capital (e.g., Holmstrom \& Kaplan 2003). However, these explanations do not predict, for example, benchmarking of the mix of components in CEO compensation. In fact, labor economics literature appears to imply that firms should choose the most efficient compensation mix among all schemes that provide the same utility to the CEO. Thus, benchmarking each pay component separately needs to be explained by new economic arguments. In this section we discuss some potential non-mutually-exclusive motivations for benchmarking pay components.

### 6.1. Extensions of the Economic Theory of Retaining CEOs

Studies such as Oyer (2005) propose that performance-based compensation can be used as a commitment device to retain employees. Vesting requirements provide incentives for employees to remain in the firm and benefit from the expected increase in the stock price. Gibbons and Murphy (1990) and De Angelis and Grinstein (2020) show that relative performance evaluation can also be used as a commitment device to retain CEOs.

Certain components of CEO compensation such as long-term incentive plans or restricted stock grants may have a stronger commitment role than salary or bonuses. Therefore, boards also consider compensation structure when incentivizing CEOs to stay in the firm, which may lead to imitating compensation structures as well as pay component levels in similar companies.

### 6.2. Unobserved CEO Preferences

To a risk averse CEO, performance-based compensation is worth less than its cost to the shareholders. The discount that CEOs apply to options or stock compensation relative to their cost to the shareholders may vary considerably depending on CEO risk aversion (Hall \& Murphy, 2002). To the extent that firms cannot fully evaluate the unique risk preferences of their CEOs, they might gravitate to a policy that provides a standard compensation mix structure, similar to the ones in other firms. Copying the pay structure of similar firms guarantees that their pay for managerial talent is competitive relative that of their competitors.

### 6.3. The Effect of Compensation Consultants

Principal-agent theories (e.g., Holmstrom, 1979) assume that the principal knows the correlation between compensation design and managerial effort and, therefore, the principal tailors the optimal compensation structure to incentivize CEO actions. However, if the principal does not know the correlation between effort, pay, and outcome, then external guidance is required. Compensation consultants usually provide this guidance for the optimal design of CEO compensation. For client firms, consultants often apply their own heuristics and understanding to information they collect regarding compensation practices in similar firms. Therefore, compensation consultants and their guiding principles may contribute to the documented attention to each pay component benchmarking and to the benchmarking of the structure of pay.

### 6.4. Asymmetric Information Between Shareholders and Public Companies

In the U.S., as well as in many other countries, shareholders are asked to ratify CEO compensation structure in their annual meetings. This Say on Pay procedure originates in concerns that boards might overcompensate CEOs or provide incentive
structures that do not align CEO and shareholder interests. Given the potential conflict of interest and asymmetric information between the board and investors regarding optimal compensation plans, investors are likely to resort to available information regarding plans in other firms to help them decide whether they should vote in favor of the given compensation plan. In addition, institutional investors typically use proxy advisory firm recommendations when casting their votes (Larcker, McCall, \& Ormazabal, 2015). These proxy advisory firms rely, at least partly, on peer-firm compensation structure data when evaluating whether a CEO compensation package is adequate. By these means, pay component benchmarking emerges.

### 6.5. Legal and Regulatory Systems

Murphy (2013) contends that the legal and regulatory systems substantially influence the design and structure of executive compensation. Disclosure rules, tax rules, accounting rules, and other regulations all push firms and directors to adopt certain compensation practices and to abandon others. A notable example is the 1994 tax rule allowing firms to deduct CEO salary up to $\$ 1$ million for tax purposes, but allowing firms to deduct any amount from the performance-based part of compensation. This ruling was shown to lead firms to keep the salary levels close to the $\$ 1$ million benchmark and also increase the weight of performance-based compensation. Rules such as the 1994 tax rule also push firms towards benchmarking CEO pay components.

For example, Regulation S-K Item 402(b)(2)(xiv) of the SEC states that the compensation committee should address in its compensation discussion and analysis section "...whether the company engaged in any benchmarking of total compensation or any material element of compensation, identifying the benchmark and, if applicable, its components, including component companies." The regulator requirements,
including the specific citing of "element" (i.e., component) of pay contribute, perhaps even directly, to the decision to benchmark pay components.

### 6.6. Director Reputation and Liabilities

Studies have shown that directors are often markedly concerned with their reputation (e.g., Jiang, Wan, \& Zhao, 2016). These concerns may lead directors to hedge potential risks of legal and social liabilities. To the extent that directors are worried that particular components of executive compensation may be inordinate and potentially launch a negative response in the media or social network channels, they may tend to align all pay components with standard levels and proportions at comparable firms.

## 7. Summary

Compensation benchmarking is an important and prevalent tool in setting CEO pay. This study proposes to examine and describe in new and extensive detail the benchmarking policies and practices of CEO pay components. We base our investigation on two samples: one, the proxy statements of S\&P 500 firms for fiscal year 2013; and two, a relatively large sample of CEO compensation data for S\&P 1500 firms (and their peers) in 2007-2013

We contribute three valuable new observations. First, we show that each major pay component is benchmarked discretely and differently from the benchmarking of total pay. Second, we establish that the adjustment of salary to that of selected peers is significantly less pronounced than the corresponding adjustments of non-equity performance pay and equity pay. Third and finally, we find that CEO's pay structure (mix of compensation components) is strongly adjusted towards that of its peer group.

A plausible interpretation of our evidence is that boards of directors recognize that each pay component has its own motivational role; consequently, it is essential to maintain a proper level of each pay component and a proper balance among all pay components. The board looks at comparable successful firms as models for prudent compensation plans. However, the case may be made that regulators, compensation consultants, proxy advisory firms, directors and shareholders all push towards pay component benchmarking. Future studies pursue investigations of these external motivations and their implications. In any case, we hope that our new evidence on pay component benchmarking will inform debates on the role and virtue of benchmarking, and will contribute to the economic understanding of CEO's pay and the pay-setting process.

## Appendix: Variables' Description

Variable Description

## I. Benchmarking related variables:

Distance in the The difference between the median weight of pay proportion of pay component X from peer group median component X in peer firms' total CEO compensation and the corresponding weight at a specific sample firm, both at year $\mathrm{t}-1$.

Ln (relative total compensation)
$\operatorname{Ln}$ (relative level of pay component X)

A benchmark measure defined as the natural logarithm of the peer-group-based total compensation target divided by firm CEO total compensation, both at year $\mathrm{t}-1$.

A benchmark measure defined as the natural logarithm of the peer-group-based target level of pay component X divided by firm CEO's level of pay component X , both at year $\mathrm{t}-1$. Sometimes abbreviated as Ln (relative pay component X) or $\operatorname{Ln}($ relative compensation component X )

## II. Compensation related variables:

All other compensation
Execucomp data item OTHCOMP, and ECA variable name OtherAnnualCompensation.

Bonus

## Equity pay

Non-equity incentive plan compensation

Non-equity performance pay

Option awards
Execucomp data item BONUS, and ECA variable name AnnualBonus. The sum of option awards and stock awards.

Execucomp data item NONEQ_INCENT, and ECA variable name NonEquityIncentivePayout.

The sum of bonus and non-equity incentive plan compensation.

Execucomp data item OPTION_AWARDS, and ECA variable name OptionAwards. For certain years (2006 in Execucomp and 2006-2008 in ECA) we use Execucomp data item OPTION_AWARDS_FV, and ECA variable name OptionGrantsISS. This facilitates consistent measurement and comparability along sample years.

The sum of change in pension value and non-qualified
Other pay
deferred compensation earnings and all other compensation.

The sum of bonus, option awards, stock awards, and nonequity incentive plan compensation.

Salary

Stock awards

Total compensation
Execucomp data item SALARY, and ECA variable name DisclosedSalary.

Execucomp data item STOCK_AWARDS, and ECA variable name StockAwards. For certain years (2006 in Execucomp and 2006-2008 in ECA) we use Execucomp data item STOCK _AWARDS_FV, and ECA variable name StockDisclosedGrantDate. This facilitates consistent measurement and comparability along sample years.

Total compensation is the sum of salary, bonus, option awards, stock awards, non-equity incentive plan compensation, change in pension value and non-qualified deferred compensation earnings, and all other compensation. Salary, bonus, option awards, stock awards, non-equity incentive plan compensation, change in pension value, and non-qualified deferred compensation earnings, and all other compensation. These compensation components disclosed in the summary compensation table of each public firm since December 2006. Execucomp data item TOTAL_SEC, and ECA variable name DisclosedTotalCompensation.

## III. Control Variables

CEO age
The age of the CEO in years.

CEO Duality
A dummy variable equal to 1 when the CEO is also the Chairman of the board (and 0 otherwise)

Total liabilities (Compustat data item LT) divided by the
Lagged leverage

Lagged Ln(sales)

Lagged Ln(monthly return standard deviation) (Compustat data items LT+CSHO*PRCC_F) at year t-1 end.

The natural logarithm of firm's sales revenue in millions of Dollars in year $\mathrm{t}-1$ (Compustat data item SALE).

The natural logarithm of the standard deviation of the monthly stock returns in the thirty-six months preceding the end of the previous fiscal year.

| Lagged market-to- <br> book value | The ratio of market value of equity to the book value of <br> equity at year t-1 end (Compustat data items <br> [CSHO*PRCC_F+TL+PSTKL-TXDITC]/AT). |
| :--- | :--- |
| ROA | Return on assets calculated as the ratio of income before <br> extraordinary items (Compustat data item IB) to total <br> assets (Compustat data item AT) in year t. |
| Stock return | The stock returns including dividends (Compustat data <br> item RET) for the current fiscal year (year t). |

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## Table 1: Company Policy Statements on Benchmarking CEO Pay

The table summarizes S\&P 500 firms' compensation benchmarking policies, as disclosed in their proxy statements (DEF 14A) for fiscal year 2013. The overall sample comprises 505 firms. We browse these firms DEF 14 A forms for information on three distinctive pay component benchmarking (salary, non-equity performance pay, and equity pay), as well as on total compensation benchmarking. The reviewed issues are presented in the form of questions with primarily "Yes/No" answers, and the number of firms in each category as well as its percentage are presented. (Percentages are rounded; thus, row total may differ slightly from $100.0 \%$.)

|  | Number of firms (\% of total in parentheses) |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Yes | No | Missing proxies or vague statements | Only total pay is benchmarked |
| Are all pay components benchmarked? | $\begin{gathered} 378 \\ (74.9 \%) \end{gathered}$ | $\begin{gathered} 75 \\ (14.9 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (5.5 \%) \end{gathered}$ | $\begin{gathered} 24 \\ (4.8 \%) \end{gathered}$ |
| Does the firm mention benchmarking only one pay component? | $\begin{gathered} 33 \\ (6.5 \%) \end{gathered}$ | $\begin{gathered} 420 \\ (83.2 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (5.5 \%) \end{gathered}$ | $\begin{gathered} 24 \\ (4.8 \%) \end{gathered}$ |
| Does the firm mention benchmarking only two pay components? | $\begin{gathered} 38 \\ (7.5 \%) \end{gathered}$ | $\begin{gathered} 415 \\ (82.2 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (5.5 \%) \end{gathered}$ | $\begin{gathered} 24 \\ (4.8 \%) \end{gathered}$ |
| Does the firm also benchmark the total pay? | $\begin{gathered} 336 \\ (66.5 \%) \end{gathered}$ | $\begin{gathered} 117 \\ (23.2 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (5.5 \%) \end{gathered}$ | $\begin{gathered} 24 \\ (4.8 \%) \end{gathered}$ |
| Does the firm benchmark at least one pay component to a specific target other than the median? | $\begin{gathered} 66 \\ (13.1 \%) \end{gathered}$ | $\begin{gathered} 387 \\ (76.6 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (5.5 \%) \end{gathered}$ | $\begin{gathered} 24 \\ (4.8 \%) \end{gathered}$ |
| Is there an indication of benchmarking the structure (mix) of pay? | $\begin{gathered} 154 \\ (30.5 \%) \end{gathered}$ | $\begin{gathered} 323 \\ (64.0 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (5.5 \%) \end{gathered}$ | - |

Table 2: Exclusion Report.
We start with all CEOs of S\&P 1500 firms in 2007-2013.

| Exclusion criteria | Number of firm-year <br> observations |
| :--- | :---: |
| Initial sample | 10,481 |
| missing compensation data | 93 |
| zero values for total compensation | 35 |
| CEOs are in their first or last year of service | 2,000 |
| The company did not specify peers | 1,830 |
| peer group comprises only 1-2 firms | 33 |
| missing compensation data for 50\% or more of the disclosed peers | 396 |
| Co-CEOs | 34 |
| firms in the financial services industry | 1,168 |
| Final sample | 4,892 |

## Table 3: Descriptive Statistics of CEO Pay and Its Components.

The sample comprises CEOs of S\&P 1500 firms in 2007-2013. Panel A reports descriptive statistics for CEO actual and peer-based target pay levels. CEO's target pay is based on peer compensation data in the previous year. All compensation figures are in thousands of dollars. Panel B reports descriptive statistics for actual and target pay changes from year t-1 to year t. Panel C reports descriptive statistics for the weight of various compensation components in total compensation at the disclosing firms, as well as the respective weights based on peer compensation data in the previous year. Target percentile is set to the median, unless the firm explicitly reports another target. All compensation figures are winsorized at the 2.5 th and 97.5 th percentiles. Definition of and details on all variables are provided in the Appendix.

Panel A: Annual pay levels of CEOs (in thousands of dollars)

|  | Mean | Std. dev. | Median | N | Mean | Std. dev. | Median | N |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Pay level among disclosing firms |  |  |  | Target pay based on peers' compensation |  |  |  |
| Total compensation | 6,964 | 5,654 | 5,260 | 4,869 | 6,850 | 4,474 | 5,617 | 4,869 |
| Salary | 884 | 320 | 850 | 4,892 | 875 | 257 | 863 | 4,892 |
| Bonus | 110 | 367 | 0 | 4,892 | 27 | 160 | 0 | 4,892 |
| Option awards | 1,258 | 1,663 | 653 | 4,891 | 1,161 | 1,159 | 836 | 4,891 |
| Stock awards | 2,343 | 2,528 | 1,500 | 4,887 | 1,936 | 1,680 | 1,475 | 4,887 |
| Non-equity incentive plan compensation | 1,296 | 1,420 | 881 | 4,891 | 1,129 | 871 | 941 | 4,891 |
| Change in pension value and nonqualified deferred compensation earnings | 524 | 1,060 | 0 | 4,876 | 365 | 696 | 0 | 4,876 |
| All other compensation | 184 | 275 | 79 | 4,891 | 140 | 136 | 103 | 4,891 |
| Aggregate pay components |  |  |  |  |  |  |  |  |
| Performance pay | 5,232 | 4,568 | 3,869 | 4,886 | 5,073 | 3,518 | 4,129 | 4,886 |
| Equity pay | 3,692 | 3,435 | 2,651 | 4,887 | 3,505 | 2,580 | 2,845 | 4,887 |
| Non-equity performance pay | 1,436 | 1,490 | 984 | 4,891 | 1,345 | 999 | 1,100 | 4,891 |
| Other pay | 778 | 1,271 | 206 | 4,697 | 617 | 825 | 263 | 4,697 |

Table 3- Continued
Panel B: Annual changes in CEO pay

|  | Mean | Std. dev. | Median | N | Mean | Std. dev. | Nedian |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Pay changes at disclosing firms |  | N |  |  |  |  |  |
| Change in Ln (total compensation ) | 0.085 | 0.44 | 0.071 | 4,859 | 0.058 | 0.24 | 0.054 |  |
| Change in Ln (salary) | 0.042 | 0.057 | 0.030 | 4,866 | 0.042 | 0.042 | 0.038 | 4,780 |
| Change in Ln (performance pay) | 0.11 | 0.57 | 0.077 | 4,769 | 0.077 | 0.31 | 0.065 | 4,696 |
| Change in Ln (equity pay) | 0.096 | 0.52 | 0.071 | 4,333 | 0.10 | 0.44 | 0.078 | 4,229 |
| Change in Ln (non-equity performance pay) | 0.037 | 0.63 | 0.038 | 4,100 | 0.025 | 0.53 | 0.020 | 4,003 |
| Change in Ln (other pay) | 0.019 | 0.83 | 0.038 | 4,697 | -0.038 | 0.57 | 0.019 | 4,621 |

Table 3-Continued

| Panel C: CEO compensation structure |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | Std. dev. | Median | N | Mean | Std. dev. | Median | N |
|  | Compensation structure at disclosing firms |  |  |  | Compensation structure at chosen peers |  |  |  |
| Salary/Total compensation | 0.19 | 0.13 | 0.16 | 4,869 | 0.17 | 0.076 | 0.15 | 4,869 |
| Bonus/Total compensation | 0.020 | 0.066 | 0 | 4,869 | 0.0036 | 0.022 | 0 | 4,869 |
| Option awards/Total compensation | 0.17 | 0.18 | 0.15 | 4,869 | 0.15 | 0.11 | 0.16 | 4,869 |
| Stock awards/Total compensation | 0.31 | 0.22 | 0.31 | 4,869 | 0.26 | 0.13 | 0.27 | 4,869 |
| Non-equity incentive plan compensation/Total compensation | 0.19 | 0.15 | 0.18 | 4,869 | 0.16 | 0.071 | 0.17 | 4,869 |
| Aggregate pay components |  |  |  |  |  |  |  |  |
| Performance pay/Total compensation | 0.71 | 0.17 | 0.75 | 4,869 | 0.73 | 0.094 | 0.75 | 4,869 |
| Equity pay/Total compensation | 0.49 | 0.21 | 0.52 | 4,869 | 0.49 | 0.13 | 0.51 | 4,869 |
| Non-equity performance pay/Total compensation | 0.21 | 0.14 | 0.19 | 4,869 |  |  |  | 4,869 |
| Other pay/Total compensation | 0.092 | 0.12 | 0.037 | 4,869 | 0.20 0.072 | 0.070 0.069 | 0.20 0.047 | 4,869 |

Table 4: Descriptive Statistics of Potential Determinants of CEO Pay.
The sample comprises CEOs of S\&P 1500 firms in 2007-2013. Stock return is the stock returns including dividends (Compustat data item RET) for the current fiscal year (year t ); ROA is the return on assets calculated as the ratio of income before extraordinary items (Compustat data item IB) to total assets (Compustat data item AT) in year t ; Lagged sales is the firm's sales revenue in millions of dollars in year $\mathrm{t}-1$ (Compustat data item SALE); Lagged monthly return standard deviation is the standard deviation of the monthly stock returns in the thirty-six months preceding the end of the previous fiscal year; Lagged market-to-book value is the ratio of market value of equity to the book value of equity at end of year t1 (Compustat data items [CSHO*PRCC_F+TL+PSTKL-TXDITC]/AT); Lagged leverage is total liabilities (Compustat data item LT) divided by the sum of total liabilities and the market value of equity (Compustat data items LT+CSHO*PRCC_F) at the end of year t-1; CEO age is the age of the CEO in years; and CEO duality is a dummy variable equal to 1 when the CEO is also the chairman of the board (and 0 otherwise).

|  | Mean | Std. dev. | Median | N |
| :--- | :---: | :---: | :---: | :---: |
| Stock return | 0.15 | 0.44 | 0.12 | 4,881 |
| ROA | 0.052 | 0.092 | 0.054 | 4,883 |
| Lagged sales (in millions of dollars) | 8,135 | 24,612 | 2,138 | 4,879 |
| Lagged monthly return standard deviation | 0.11 | 0.050 | 0.10 | 4,641 |
| Lagged market-to-book value | 1.82 | 1.08 | 1.51 | 4,501 |
| Lagged leverage | 0.35 | 0.20 | 0.33 | 4,867 |
| CEO age | 56 | 7 | 56 | 4,883 |
| CEO duality | 0.16 | 0.36 | 0 | 4,892 |

## Table 5: A Baseline Model of the Level of CEO Compensation.

The table reports regression estimates of Equation 1. The sample comprises CEOs of S\&P 1500 firms in 2007-2013. The dependent variables are the natural logarithm of CEO's total compensation, salary, performance pay, non-equity performance pay, equity pay, stock awards and option awards, all in year t . Definition of and details on all variables are provided in the Appendix. Year $\times$ Industry FE are dummy variables for each unique combination of industry and year based on Fama and French (1997)'s 49-industry classification. Standard errors, reported in parentheses, are clustered at the firm level. ${ }^{* * *}$, ${ }^{* *}$, and $*$ denote significance at the $1 \%, 5 \%$, and $10 \%$ levels, respectively.

## $\mathrm{Ln}(\mathrm{CEO}$ compensation)

|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total compensation | Salary | Performance pay | Non-equity performance pay | Equity pay | Stock awards | Option awards |
| Intercept | $\begin{gathered} 4.80^{* * *} \\ (0.23) \end{gathered}$ | $\begin{gathered} 5.15 * * * \\ (0.10) \end{gathered}$ | $\begin{gathered} 4.48^{* * *} \\ (0.27) \end{gathered}$ | $\begin{gathered} 3.027 * * * \\ (0.27) \end{gathered}$ | $\begin{gathered} 4.27 * * * \\ (0.33) \end{gathered}$ | $\begin{gathered} 4.32 * * * \\ (0.30) \end{gathered}$ | $\begin{gathered} 3.59 * * * \\ (0.39) \end{gathered}$ |
| Lagged Ln(sales) | $\begin{aligned} & 0.40^{* * *} \\ & (0.017) \end{aligned}$ | $\begin{aligned} & 0.17^{* * *} \\ & (0.0078) \end{aligned}$ | $\begin{aligned} & 0.45 * * * \\ & (0.022) \end{aligned}$ | $\begin{aligned} & 0.37 * * * \\ & (0.023) \end{aligned}$ | $\begin{gathered} 0.50 * * * \\ (0.018) \end{gathered}$ | $\begin{gathered} 0.44 * * * \\ (0.021) \end{gathered}$ | $\begin{aligned} & 0.47 * * * \\ & (0.021) \end{aligned}$ |
| Stock return | $\begin{gathered} 0.22 * * * \\ (0.026) \end{gathered}$ | $\begin{aligned} & -0.0058 \\ & (0.010) \end{aligned}$ | $\begin{gathered} 0.30^{* * *} \\ (0.035) \end{gathered}$ | $\begin{gathered} 0.47 * * * \\ (0.040) \end{gathered}$ | $\begin{gathered} 0.13^{* * *} \\ (0.036) \end{gathered}$ | $\begin{aligned} & 0.10^{* *} \\ & (0.049) \end{aligned}$ | $\begin{gathered} 0.18^{* * *} \\ (0.040) \end{gathered}$ |
| One-year lagged stock return | $\begin{gathered} 0.16^{* * *} \\ (0.026) \end{gathered}$ | $\begin{aligned} & 0.0086 \\ & (0.010) \end{aligned}$ | $\begin{gathered} 0.23 * * * \\ (0.035) \end{gathered}$ | $\begin{aligned} & 0.22 * * * \\ & (0.037) \end{aligned}$ | $\begin{aligned} & 0.11 * * * \\ & (0.034) \end{aligned}$ | $\begin{gathered} 0.13 * * * \\ (0.039) \end{gathered}$ | $\begin{aligned} & 0.10^{* *} \\ & (0.038) \end{aligned}$ |
| ROA | $\begin{aligned} & 0.030 \\ & (0.12) \end{aligned}$ | $\begin{aligned} & 0.12^{* *} \\ & (0.057) \end{aligned}$ | $\begin{gathered} 0.15 \\ (0.18) \end{gathered}$ | $\begin{gathered} 1.51^{* * *} \\ (0.29) \end{gathered}$ | $\begin{gathered} -0.38 * * \\ (0.18) \end{gathered}$ | $\begin{gathered} -0.092 \\ (0.23) \end{gathered}$ | $\begin{gathered} -0.66^{* * *} \\ (0.20) \end{gathered}$ |
| One-year lagged ROA | $\begin{gathered} -0.32 * * \\ (0.14) \end{gathered}$ | $\begin{aligned} & -0.14^{* *} \\ & (0.059) \end{aligned}$ | $\begin{aligned} & -0.42^{*} \\ & (0.22) \end{aligned}$ | $\begin{aligned} & -0.46^{*} \\ & (0.27) \end{aligned}$ | $\begin{gathered} -0.49^{* *} \\ (0.19) \end{gathered}$ | $\begin{gathered} -0.37^{*} \\ (0.21) \end{gathered}$ | $\begin{gathered} -0.62 * * * \\ (0.22) \end{gathered}$ |

Table 5-Continued

|  | Ln (CEO compensation) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
|  | Total compensation | Salary | Performance pay | Non-equity performance pay | Equity pay | Stock awards | Option awards |
| Lagged $\operatorname{Ln}$ (monthly return standard deviation) | -0.14*** | -0.017 | -0.14** | -0.10 | -0.081 | -0.12 | 0.10 |
|  | (0.047) | (0.021) | (0.061) | (0.070) | (0.060) | (0.075) | (0.073) |
| Lagged market-to-book value | 0.045** | -0.016 | 0.044 | 0.0016 | 0.12*** | 0.052 | 0.20 *** |
|  | (0.021) | (0.010) | (0.029) | (0.026) | (0.024) | (0.032) | (0.029) |
| Lagged leverage | -0.14 | 0.036 | -0.27* | 0.0032 | $-0.53 * * *$ | -0.38** | -0.71 *** |
|  | (0.12) | (0.057) | (0.15) | (0.18) | (0.17) | (0.19) | (0.19) |
| CEO Age | 0.0030 | 0.0049*** | -0.0035 | 0.011*** | -0.0058* | -0.0040 | -0.0039 |
|  | (0.0024) | (0.0013) | (0.0031) | (0.0033) | (0.0034) | $(0.0041)$ | (0.0041) |
| CEO Duality | 0.19*** | 0.10*** | 0.16*** | 0.21*** | 0.11* | 0.10 | 0.15** |
|  | (0.044) | (0.018) | (0.059) | (0.063) | (0.062) | (0.079) | (0.067) |
| Year $\times$ Industry FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 4,255 | 4,275 | 4,269 | 3,839 | 4,035 | 3,629 | 2,701 |
| Adjusted $\mathrm{R}^{2}$ | 0.56 | 0.62 | 0.47 | 0.41 | 0.48 | 0.38 | 0.44 |

## Table 6: Preliminary Evidence on Benchmarking in CEO Compensation and Its Components.

The sample comprises CEOs of S\&P 1500 firms in 2007-2013. Panel A compares changes to CEO pay from year t - 1 to year t based on CEO's pay status relative to peers in the previous year. The table shows the mean and the median logarithmic changes in pay for CEOs who are paid above the peer-based target pay and for CEOs who are paid below the peer-based target pay in the previous year. The analysis examines nine forms of CEO pay as follows: total compensation, total compensation excluding the change in pension value, salary, performance pay, non-equity performance pay, equity pay, stock awards, option awards, and other pay. The Wilcoxon signed rank-sum test and t-test are used to assess statistical significance for differences in median and mean, respectively, between the above and below the target groups. Panel B focuses on the changes in the weight of various pay components in total compensation from year $t-1$ to year $t$. It compares two subsamples: 1) CEOs whose pay components weight in total compensation was above the peer group median in the previous year; and 2) CEOs whose pay component weight in total compensation was below the peer group median in the previous year. A $t$-test is used to assess the statistical significance of weight changes between the above and below target groups. Panel C presents the results of a regression that examines benchmarking in the structure of CEO pay. The dependent variable is the change in the proportion of the pay component from year $\mathrm{t}-1$ to year t , while the explanatory variable is the year $\mathrm{t}-1$ distance of pay component weight in total compensation from the median respective weight amongst peer CEOs. Year $\times$ Industry FE are dummy variables for each unique combination of industry and year, where industry is based on Fama and French (1997)'s 49-industry classification. Standard errors, reported in parentheses, are clustered at the firm level. ${ }^{* * *}$, ${ }^{* *}$, and $*$ denote significance at the $1 \%, 5 \%$, and $10 \%$ levels, respectively. Definitions and details for all variables are provided in the Appendix.

Table 6-Continued

| Panel A |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Pay measure | Group | Number of observations | Mean change in pay | Median change in pay | $p$-Values for difference (one sided test) |  |
|  |  |  |  |  | $t$-Test | Wilcoxon test |
| Total compensation | Above target | 1,853 | -0.087 | -0.027 | <. 0001 | <. 0001 |
|  | Below target | 3,006 | 0.19 | 0.14 |  |  |
| Total compensation (excluding change in pension value) | Above target | 1,873 | -0.087 | -0.020 | <. 0001 | <. 0001 |
|  | Below target | 3,019 | 0.20 | 0.14 |  |  |
| Salary | Above target | 2,012 | 0.025 | 0.020 | <. 0001 | <. 0001 |
|  | Below target | 2,854 | 0.053 | 0.039 |  |  |
| Performance Pay | Above target | 1,909 | -0.11 | -0.027 | <. 0001 | <. 0001 |
|  | Below target | 2,860 | 0.25 | 0.18 |  |  |
| Non-equity performance pay | Above target | 2,077 | -0.12 | -0.041 | <. 0001 | <. 0001 |
|  | Below target | 2,023 | 0.19 | 0.15 |  |  |
| Equity pay | Above target | 1,934 | -0.090 | -0.0040 | <. 0001 | <. 0001 |
|  | Below target | 2,399 | 0.25 | 0.17 |  |  |
| Stock awards | Above target | 2,023 | -0.030 | 0.026 | <. 0001 | <. 0001 |
|  | Below target | 1,659 | 0.32 | 0.19 |  |  |
| Option awards | Above target | 1,891 | -0.064 | -0.000020 | <. 0001 | <. 0001 |
|  | Below target | 899 | 0.19 | 0.11 |  |  |
| Other pay | Above target | 2,060 | -0.16 | -0.0039 | <. 0001 | <. 0001 |
|  | Below target | 2,637 | 0.16 | 0.067 |  |  |

## Table 6-Continued

Panel B: Preliminary evidence on the benchmarking of the structure of CEO pay

| Pay component | Group | Number of observations | Mean weight of pay component in total compensation in year t-1 | Mean change in the weight of the pay component | $p$-value of the change (based on a one-sided $t$ - test) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Salary | Above median | 2,798 | 0.26 | -0.037 | <. 0001 |
|  | Below median | 2,061 | 0.13 | 0.025 |  |
| Performance pay | Above median | 2,409 | 0.80 | -0.036 | <. 0001 |
|  | Below median | 2,450 | 0.59 | 0.063 |  |
| Non-equity performance pay | Above median | 2,555 | 0.32 | -0.073 | <. 0001 |
|  | Below median | 2,304 | 0.12 | 0.051 |  |
| Equity pay | Above median | 2,357 | 0.62 | -0.053 | <. 0001 |
|  | Below median | 2,502 | 0.32 | 0.10 |  |
| Stock awards | Above median | 2,403 | 0.44 | -0.031 | <. 0001 |
|  | Below median | 2,456 | 0.13 | 0.094 |  |
| Option awards | Above median | 2,269 | 0.33 | -0.059 | <. 0001 |
|  | Below median | 2,590 | 0.050 | 0.042 |  |
| Other pay | Above median | 2,322 | 0.16 | -0.022 | <. 0001 |
|  | Below median | 2,537 | 0.033 | 0.014 |  |

Table 6-Continued
Panel C: Regression tests of benchmarking in the structure of CEO pay
Change in the weight of pay component X in total compensation

|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | (6) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Salary | Performance <br> pay | Non-equity <br> performance <br> pay | Equity pay | Stock awards | Option <br> awards |
| Distance from peer group's median weight | $0.50^{* * *}$ | $0.49^{* * *}$ | $0.54 * * *$ | $0.5^{* * *}$ | $0.34^{* * *}$ | $0.34 * * *$ |
| Year $\times$ Industry FE | $(0.027)$ | $(0.020)$ | $(0.02)$ | $(0.02)$ | $(0.015)$ | $(0.016)$ |
| Observations | Yes | Yes | Yes | Yes | Yes | Yes |
| Adjusted $\mathrm{R}^{2}$ | 4,859 | 4,859 | 4,859 | 4,859 | 4,859 | 4,859 |

Table 7: The Effect of Benchmarking on the Yearly Revision in CEO Pay Components.
The table presents the results of fitting Equation 4. The sample comprises CEOs of S\&P 1500 firms in 2007-2013. Definition of and details on all variables are provided in the Appendix. Year $\times$ Industry FE are dummy variables for each unique combination of industry and year, using the 49industry classification of Fama and French (1997). Note that for each pay component, we fit an individual parsimonious model that is restricted to include only explanatory variables that are significant at the $1 \%$ level at least in our basic pay components regressions (see Table 5). Further, statistically insignificant coefficients are omitted from the table. Standard errors, reported in parentheses, are clustered at the firm level. ${ }^{* * *}$, **, and $*$ denote significance at the $1 \%, 5 \%$, and $10 \%$ levels, respectively.

Change in Ln (CEO compensation component X )


Table 7-Continued

|  | Change in Ln (CEO compensation component X ) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) |
|  | Salary | Performance pay | Non-equity performance pay | Equity pay | Stock awards | Option awards |
| Change in one-year lagged stock return |  | $\begin{gathered} 0.086^{* * *} \\ (0.031) \end{gathered}$ | $\begin{gathered} 0.092^{* *} \\ (0.043) \end{gathered}$ | $\begin{gathered} 0.028^{* *} \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.055 * * * \\ (0.017) \end{gathered}$ |  |
| Change in ROA |  |  | $\begin{gathered} 1.27^{* * *} \\ (0.25) \end{gathered}$ |  |  |  |
| Change in lagged ROA |  |  |  |  |  |  |
| Change in lagged $\operatorname{Ln}$ (monthly return standard deviation) |  |  |  |  |  |  |
| Change in lagged market-to-book value |  |  |  | $\begin{gathered} 0.064 * * * \\ (0.018) \end{gathered}$ |  | $\begin{gathered} 0.069 * * * \\ (0.025) \end{gathered}$ |
| Change in lagged leverage |  |  |  | $\begin{gathered} -0.49 * * * \\ (0.14) \end{gathered}$ |  | $\begin{gathered} -0.68^{* * *} \\ (0.21) \end{gathered}$ |
| Year $\times$ Industry FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 4,828 | 4,683 | 3,997 | 3,822 | 3,467 | 2,183 |
| Adjusted R ${ }^{2}$ | 0.18 | 0.31 | 0.25 | 0.22 | 0.18 | 0.12 |

Table 8: Variation in Benchmarking Across the Three Central Pay Components.
Panel A presents the results of fitting Equation 4 on a system of three key pay components (salary, non-equity performance pay, and equity pay) using seemingly unrelated regressions. Panel B reports F-tests of the differences in benchmarking coefficients across our three pay components. The sample comprises CEOs of S\&P 1500 firms in 2007-2013. Definition of and details on all variables are provided in the Appendix. Year $\times$ Industry FE are dummy variables for each unique combination of industry and year based on the 49 -industry classification of Fama and French (1997). Note that for each pay component, we employ an individual parsimonious model that is restricted to include only explanatory variables that are significant at the $1 \%$ level at least in our basic pay components regressions (see Table 5). Standard errors are reported in parentheses. ***, ${ }^{* *}$, and $*$ denote significance at the $1 \%, 5 \%$, and $10 \%$ levels, respectively.

## Panel A

Results from fitting Equation 5 using seemingly unrelated regressions.

|  | Change in Ln (CEO compensation component X ) |  |  |
| :---: | :---: | :---: | :---: |
|  | (1) <br> Salary | (2) Non-equity performance pay | (3) <br> Equity pay |
| Intercept | 0.056*** | 0.084** | 0.21 *** |
|  | (0.0030) | (0.033) | (0.028) |
| Ln (relative level of pay component X ) | 0.089*** | 0.22*** | 0.19*** |
|  | (0.0045) | (0.017) | (0.015) |
| Distance of pay component X weight from peer group's median weight | $\begin{aligned} & 0.028^{* *} \\ & (0.012) \end{aligned}$ | $\begin{gathered} 0.37 * * * \\ (0.11) \end{gathered}$ | $0.54 * * *$ (0.072) |
| Other explanatory variables as in Table 7 | Yes | Yes | Yes |
| Year $\times$ Industry FE | Yes | Yes | Yes |
| Observations | 3,244 | 3,244 | 3,244 |
| System Weighted $\mathrm{R}^{2}$ |  | 0.285 |  |

Table 8-Continued

| Panel B <br> Examining differences in benchmarking across pay components. |  |  |
| :--- | :---: | :---: |
| H0: The coefficients of Ln(relative level of pay <br> component X) are equal in the equations of | $F$-statistic | $p$-value |
| Salary, non-equity performance pay and equity <br> pay | 44.33 | 0.0001 |
| Non-equity performance pay and equity pay | 2.14 | 0.14 |
|  |  |  |
| H0: The coefficients of Distance from peer <br> group's median weight are equal in the equations <br> of | $F$-statistic | $p$-value |
| Salary, non-equity performance pay and equity <br> pay | 29.54 | 0.0001 |
| Non-equity performance pay and equity pay | 1.65 | 0.20 |

Table 9: Tests of the Difference in Benchmarking Between Total Compensation and Pay Components.
The table presents the results of fitting equation (6) to a system of three key pay components (salary, non-equity performance pay, and equity pay) using seemingly unrelated regressions. The sample comprises CEOs of S\&P 1500 firms in 2007-2013. Definition of and details on all variables are provided in the Appendix. Year $\times$ Industry FE are dummy variables for each unique combination of industry and year based on the 49-industry classification of Fama and French (1997). Note that for each pay component, we employ an individual parsimonious model that is restricted to include only explanatory variables that are significant at the $1 \%$ level at least in our basic pay components regressions (see Table 5). To overcome multicollinearity between relative total compensation and relative pay component $X$, we first regress each relative pay component $X$ on relative total compensation. Then, we use the residuals of these regressions instead of the relative pay components in the regressions. Standard errors are reported in parentheses. ${ }^{* * *},{ }^{* *}$, and $*$ denote significance at the $1 \%, 5 \%$, and $10 \%$ levels, respectively.

Table 9- Continued

|  | Change in Ln (CEO compensation component X ) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) |
|  | Salary | Non-equity performance pay | Equity pay | Salary | Non-equity performance pay | Equity pay |
| Intercept | 0.057*** | $0.12 * * *$ | 0.14*** |  | 0.082** | 0.19*** |
|  | (0.0031) | (0.034) | (0.028) | (0.0030) | (0.033) | (0.029) |
| $\operatorname{Ln}$ (relative total compensation) | $\begin{gathered} 0.015^{* * *} \\ (0.0020) \end{gathered}$ | $\begin{aligned} & 0.20 * * * \\ & (0.022) \end{aligned}$ | $\begin{gathered} 0.37 * * * \\ (0.018) \end{gathered}$ |  |  |  |
| $\operatorname{Ln}$ (relative level of pay component X ) |  |  |  | $\begin{gathered} 0.089 * * * \\ (0.0045) \end{gathered}$ | $\begin{gathered} 0.26^{* * *} \\ (0.013) \end{gathered}$ | $\begin{gathered} 0.26 * * * \\ (0.011) \end{gathered}$ |
| Other explanatory variables as in Table 7 | Yes | Yes | Yes | Yes | Yes | Yes |
| Year $\times$ Industry FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 3,285 | 3,285 | 3,285 | 3,250 | 3,250 | 3,250 |
| System Weighted $\mathrm{R}^{2}$ |  | 0.2192 |  |  | 0.2805 |  |

Table 9- Continued

|  | Change in Ln (CEO compensation component X ) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (7) | (8) | (9) | (10) | (11) | (12) |
|  | Salary | Non-equity performance pay | Equity pay | Salary | Non-equity performance pay | Equity pay |
| Intercept | 0.058*** | 0.064* | 0.17*** | 0.056*** | 0.084** | 0.21*** |
|  | (0.0030) | (0.033) | (0.029) | (0.0030) | (0.033) | (0.028) |
| $\operatorname{Ln}$ (relative total compensation) | 0.012*** | 0.19*** | 0.38*** |  |  |  |
|  | (0.0019) | (0.021) | (0.018) |  |  |  |
| $\operatorname{Ln}$ (relative level of pay component X ) | 0.094*** | 0.26*** | 0.20*** | 0.089*** | 0.22*** | 0.19*** |
|  | (0.0051) | (0.014) | (0.016) | (0.0045) | (0.017) | (0.015) |
| Distance from peer group's median weight |  |  |  |  |  |  |
|  |  |  |  | (0.012) | (0.11) | (0.072) |
| Other explanatory variables as in Table 7 | Yes | Yes | Yes | Yes | Yes | Yes |
| Year $\times$ Industry FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 3,244 | 3,244 | 3,244 | 3,244 | 3,244 | 3,244 |
| System Weighted $\mathrm{R}^{2}$ |  | 0.2820 |  |  | 0.2854 |  |


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[^1]:    ${ }^{1}$ We note that benchmarking each component does not necessarily imply that the mix itself is also benchmarked. For example, data from compensation committee reports show that benchmarking components to the median levels of peers is a common practice. For most distributions, this finding does not imply that the ratio of each component out of total compensation is also benchmarked to the median ratio of peers.

[^2]:    ${ }^{2}$ A growing strand of literature provides evidence for the role and the effect of peer firms beyond compensation benchmarking. Peer selection also affects relative performance awards (RPE)-see, e.g., Bizjak, Kalpathy, Li, and Young (2020); De Angelis and Grinstein (2020); and Ma, Shin, and Wang (2019). Peer groups also play an important role in other corporate policies such as corporate investment, corporate capital structure and financial policies (e.g., Foucault \& Fresard, 2014; Leary \& Roberts, 2014).

[^3]:    ${ }^{3}$ We are grateful to Ana Albuquerque and her coauthors for providing us with these data.

[^4]:    ${ }^{4}$ Peer CEO pay is marked as missing also in years when the peer CEO was replaced or appointed.

[^5]:    ${ }^{5}$ Among pay components, base salary appears the most distinct. CEO base salary is significantly affected only by firm size, CEO age, and the CEO-chairman duality.

[^6]:    ${ }^{6}$ We also run a set of regressions with three benchmarking variables: total pay; pay component; and pay mix benchmarking. However, these regressions are plugged with severe multicollinearity problems that obstruct any inference regarding any single benchmarking variable.

[^7]:    ${ }^{7}$ In 2009, the SEC adopted amendments requiring reporting of the aggregate grant date fair value of stock and option awards in the summary compensation table, in accordance with FASB ASC Topic 718, instead of reporting stock and option awards at fair value measured by the amount expensed under FAS 123(R). Pre-2009 data comparable to post-2009 measurement of stock and option awards are reported in the plan-based awards table of the compensation report. Accordingly, for pre-2009 years, we use the values from the plan-based awards table.
    ${ }^{8}$ For example, the coefficient of $\operatorname{Ln}\left(\right.$ Relative pay component $\left._{i, t-1}\right)$ in the performance pay regression decreases from 0.37 in the first subperiod to 0.29 in the second.

